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ABSTRACT

This paper addresses the positive implications of indexing risky debt to observable aggregate conditions. These issues are pursued within the context of the celebrated financial accelerator model of Bernanke, Gertler and Gilchrist (1999). The principle conclusions include: (1) the estimated level of indexation is significant, (2) the business cycle properties of the model are significantly affected by this degree of indexation, (3) the importance of investment shocks in the business cycle depends upon the estimated level of indexation, and (4) although the data prefers the financial model with indexation over the frictionless model, they have remarkably similar business cycle properties for non-financial exogenous shocks.

JEL Codes: E32, E44.

Keywords: Agency costs; financial accelerator; business cycles.
1. INTRODUCTION

The fundamental function of credit markets is to channel funds from savers to entrepreneurs who have some valuable capital investment project. These efforts are hindered by agency costs arising from asymmetric information. A standard result in a subset of this literature, the costly state verification (CSV) framework, is that risky debt is the optimal contract between risk-neutral lenders and entrepreneurs. The modifier risky simply means that there is a non-zero chance of default. In the CSV model, external parties can observe the realization of the entrepreneur’s idiosyncratic production technology only by expending a monitoring cost. Townsend (1979) demonstrates that risky debt is optimal in this environment because it minimizes the need for verification of project outcomes. This verification is costly but necessary to align the incentives of the firm with the bank.

Aggregate conditions will also affect the ability of the borrower to repay the loan. But since aggregate variables are observed by both parties, it may be advantageous to have the loan contract indexed to the behavior of aggregate variables. Therefore, even when loan contracts cannot be designed based on private information, we can exploit common information to make these financial contracts more state-contingent. That is, why should the loan contract call for costly monitoring when the event that leads to a poor return is observable by all parties? Carstom, Fuerst, and Paustian (2013) examine questions of this type within the financial accelerator of Bernanke, Gertler, and Gilchrist (1999), hereafter BGG. Carstom et al. (2013) demonstrate that the privately optimal contract in the BGG model includes indexation to: \( i \) the aggregate return to capital (which we will call \( R^k \)-indexation), \( ii \) the marginal utility of wealth (which we will call \( \lambda \)-indexation), and \( iii \) the shadow cost of external financing.

\[ ^1 \text{This is the logic behind Shiller and Weiss’s (1999) suggestion of indexing home mortgages to movements in aggregate house prices.} \]
In this paper we explore the business cycle implications of indexing the BGG loan contract to the aggregate return to capital and to the marginal utility of wealth. There are at least two reasons why this is an interesting exercise. First, as noted above, Carlstrom et al. (2013) demonstrate that the privately optimal contract in the BGG framework includes indexation of this very type. Second, indexation of this type is not so far removed from some financial contracts we do observe. For example, indexing repayment to innovations in the marginal utility of wealth is a close approximation to indexation to movements in the risk free rate of interest. There are many debt instruments that are directly linked to market interest rate of this type, e.g., adjustable rate mortgages. More generally, since we are assuming that the CSV framework proxies for agency cost effects in the entire US financial system, it seems reasonable to include some form of indexation to mimic the myriad ex post returns on external financing. For example, in contrast to the model assumption where entrepreneurs get zero in the event of bankruptcy, this is clearly not the implication of Chapter 11 bankruptcy. In any event, we use familiar Bayesian methods to estimate the degree of contract indexation to the return to capital and the marginal utility of wealth.

To avoid misspecification problems in the estimation we need a complete model of the business cycle. We use the recent contribution of Justiniano, Primiceri, and Tambalotti (2011), hereafter JPT, as our benchmark. A novelty of the JPT model is that it includes two shocks to the capital accumulation technology. The first shock is a non-stationary shock to the relative cost of producing investment goods, the “investment specific technology shock” (IST). The second is a stationary shock to the transformation of investment goods into installed capital, the “marginal efficiency of investment shock” (MEI). For business cycle variability, JPT find that the IST shocks are irrelevant, while the MEI shocks account for a substantial portion of business cycle fluctuations.

Our principle results include the following. First, the estimated level of \( R^k \)-indexation significantly exceeds unity, much higher than the assumed BGG indexation of approximately zero. A model with \( R^k \)-indexation fits the data significantly better when compared to BGG. This is because the BGG model’s prediction for the risk premium in
the wake of a MEI shock is counterfactual. A MEI shock lowers the price of capital and thus leads to a sharp decline in entrepreneurial net worth in the BGG model. But under $R^k$-indexation, the required repayment falls also so that net worth moves by significantly less.

Second, with $R^k$-indexation, this financial model and JPT have remarkably similar business cycle properties for non-financial exogenous shocks. For example, for the case of MEI shocks, the estimated level of indexation leads to net worth movements in the financial model that accommodate real behavior quite similar to the response of JPT to a MEI shock. We also nest financial shocks into the JPT model by treating fluctuations in these two financial variables as serially correlated measurement error. This model horse race results in the $R^k$-indexation model dominating BGG, which in turn significantly dominates JPT. The financial models are improvements over JPT in two ways. The financial models make predictions for the risk premium and leverage on which JPT is silent, and the financial models introduce other exogenous shocks, e.g., shocks to net worth or idiosyncratic variance, that are irrelevant in JPT.

Third, we find that whether financial shocks or MEI shocks are more important drivers of the business cycle depends upon the level of indexation. Under BGG, financial shocks account for a significant part of the variance of investment spending. But under the estimated level of $R^k$-indexation, financial shocks become much less important and the MEI shocks are again of paramount importance.

Two prominent papers closely related to the current work are Christiano, Motto, and Rostagno (2010), and DeGraeve (2008). They each use Bayesian methods to estimate versions of the BGG framework in medium-scale macro models. Both papers conclude that the model with financial frictions provides a better fit to the data when compared to its frictionless counterpart. The chief novelty of the current paper is to introduce contract indexation into the BGG framework, and demonstrate that it is empirically relevant, altering the business cycle properties of the model. Neither of the previous papers considered indexation of this type.
The paper proceeds as follows. Section 2 presents a simple example that illustrates the importance of contract indexation to the financial accelerator. Section 3 develops the DSGE model. Section 4 presents the estimation results. Section 5 concludes.

2. WHY DOES INDEXATION MATTER? A SIMPLE EXAMPLE

This section presents a simple intuitive example that demonstrates the importance of indexation in determining the size of the financial accelerator. Consider a world with agency costs in which the portion of net worth owned by entrepreneurs \( nw_i \) has a positive effect on the value of capital \( q_i \):

\[
q_i = p^*nw_i + \epsilon^d_i
\]  

(1)

where the expression is in log deviations and \( \epsilon^d_i \) is an exogenous shock to capital prices, e.g., a shock to MEI in the general equilibrium model below. Equation 1 is a manifestation of agency costs in that the distribution of net worth across lenders and borrowers affects asset prices. The idea is that higher net worth in the hands of entrepreneurs makes it easier for them to access a loan with which to buy capital, so that higher levels of net worth act like a demand channel on asset prices. In the general equilibrium model below, the value of \( p \) is a function of the agency cost and (installed) capital adjustment cost parameters.

The entrepreneur accumulates net worth to mitigate the agency problems involved in direct lending. The agency problem arises from a CSV problem in the entrepreneur’s production technology. The entrepreneur takes one unit of input and creates \( \omega_i \) units of capital, where the unit-mean random variable \( \omega_i \) is privately observed by the entrepreneur but can be verified by the lender only by paying a cost. This CSV problem makes equity finance problematic, so that the optimal contract is given by a risky debt contract with a promised repayment of \( r^p_i \). The repayment \( r^p_i \) cannot be indexed to \( \omega_i \) because it is privately observed. But it can be indexed to the aggregate price of capital:

\[
r^p_i = \chi q_i.
\]  

(2)
This form of indexation is similar to indexing to the rate of return on capital in the general equilibrium model developed below.

Entrepreneurial net worth accumulates with the profit flow from the investment project, but decays via consumption of entrepreneurs (which is a constant fraction of net worth). Log-linearized this evolution is given by:

\[ nw_t = \kappa (q_t - r^n_t) + nw_{t-1} + r^n_t + \epsilon^n_t \]  

where \( \kappa > 1 \) denotes leverage (the ratio of project size to net worth) and \( \epsilon^n_t \) is an exogenous shock to net worth. Using the indexation assumption (2), we can express (3) as

\[ nw_t = q_t [\kappa - \chi (\kappa - 1)] + nw_{t-1} + \epsilon^n_t \]  

Note that since \( \kappa > 1 \), the slope of the net worth equation is decreasing in the level of indexation.

Equations 1 and 4 are a simultaneous system in net worth and the price of capital. We can solve for the two endogenous variables as a function of the pre-determined and exogenous variables:

\[ nw_t = \frac{nw_{t-1} + \epsilon^n_t + \epsilon^d_t [\kappa - \chi (\kappa - 1)]}{1 - p[\kappa - \chi (\kappa - 1)]} \]  

\[ q_t = \frac{p(nw_{t-1} + \epsilon^n_t) + \epsilon^d_t}{1 - p[\kappa - \chi (\kappa - 1)]} \]  

The inverse of the denominator in Equations 5-6 is the familiar multiplier arising from two endogenous variables with positive feedback. This then implies that exogenous shocks are multiplied or financially accelerated, and that the degree of this multiplication depends upon the level of indexation. The effect of indexation on the financial multiplier is highly nonlinear. Figure 1 plots the multiplier for \( \kappa = 2 \), and \( p = 0.45 \), both of these values roughly correspond to the general equilibrium analysis below. Note that moving from \( \chi = 0 \) to \( \chi = 1 \), has an enormous effect on the multiplier. But there are sharp diminishing returns so the multiplier is little changed as we move from \( \chi = 1 \) to \( \chi = 2 \). This suggests, and we confirm below, that the data can distinguish \( \chi = 0 \) from, say, \( \chi = 1 \),
but that indexation values in excess of unity will have similar business cycle characteristics and thus be difficult to identify.

Consider three special cases of indexation: $\chi = 0$, $\chi = 1$, and $\chi = \frac{\kappa}{(\kappa - 1)}$. The first is the implicit assumption in BGG; the second implies complete indexation; the third eliminates the financial accelerator altogether. In these cases, net worth and asset prices are given by:

<table>
<thead>
<tr>
<th>Indexation</th>
<th>Net worth</th>
<th>Capital price</th>
<th>Multiplier (p=0.45, $\kappa = 2$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\chi = 0$</td>
<td>$\frac{nw_{i-1} + \epsilon_i^n + \kappa \epsilon_i^d}{1 - p \kappa}$</td>
<td>$\frac{p(nw_{i-1} + \epsilon_i^n) + \epsilon_i^d}{1 - p \kappa}$</td>
<td>10</td>
</tr>
<tr>
<td>$\chi = 1$</td>
<td>$\frac{nw_{i-1} + \epsilon_i^n + \epsilon_i^d}{1 - p}$</td>
<td>$\frac{p(nw_{i-1} + \epsilon_i^n) + \epsilon_i^d}{1 - p}$</td>
<td>1.82</td>
</tr>
<tr>
<td>$\chi = \frac{\kappa}{(\kappa - 1)}$</td>
<td>$nw_{i-1} + \epsilon_i^n$</td>
<td>$p(nw_{i-1} + \epsilon_i^n) + \epsilon_i^d$</td>
<td>1</td>
</tr>
</tbody>
</table>

For both $\chi = 0$ and $\chi = 1$, exogenous shocks to asset prices and net worth have multiple effects on the equilibrium levels of net worth and capital prices. Since $\kappa > 1$, this effect is much larger under BGG’s assumption of no indexation ($\frac{1}{1 - p \kappa} \gg \frac{1}{1 - p}$).

Further, under the BGG assumption, exogenous shocks to asset prices ($\epsilon_i^d$) have an added effect as they are weighted by leverage. But for all levels of indexation, there are always agency cost effects in that the price of capital is affected by the level of entrepreneurial net worth. The financial multiplier effects are traced out in Figure 2: an exogenous shock to asset prices has a much larger effect on both net worth and asset prices in the BGG framework. Finally, since $\kappa \approx 2$ the financial accelerator largely disappears when $\chi = 2$.

Before proceeding, it is helpful to emphasize the two parameters that are crucial in our simple example as they will be manifested below in the richer general equilibrium.
environment. Our reduced form parameter \( p \) in Equation 1 is the agency cost parameter. In a Modigliani-Miller world we would have \( p = 0 \), as the distribution of net worth would have no effect on asset prices or real activity. Second, the indexation parameter \( \chi \) determines the size of the financial accelerator, i.e., how do unexpected movements in asset prices feed in to net worth? These are two related but logically distinct ideas. That is, one can imagine a world with agency costs \( (p > 0) \), but with very modest accelerator effects \( (\chi = \frac{\kappa}{(k-1)}) \). To anticipate our empirical results, this is the parameter set that wins the model horse race. That is, the data is consistent with an agency cost model but with trivial accelerator effects. In such an environment, financial shocks (e.g., shocks to net worth) will affect real activity, but other real shocks (e.g., MEI shocks) will not be accelerated.

3. THE MODEL

The benchmark model follows the JPT framework closely. The model of agency costs comes from BGG with the addition of exogenous contract indexation. The BGG loan contract is between lenders and entrepreneurs, so we focus on these two agents first before turning to the familiar framework of JPT.

LENDERS

The representative lender accepts deposits from households (promising a sure real return \( R^d_t \)) and provides loans to the continuum of entrepreneurs. These loans are intertemporal, with the loans made at the end of time \( t \) being paid back in time \( t+1 \). The realized gross real return on these loans is denoted by \( R^L_{t+1} \). Each individual loan is subject to idiosyncratic and aggregate risk, but since the lender holds an entire portfolio of loans, only the aggregate risk remains. The lender has no other source of funds, so the level of loans will equal the level of deposits. Hence, real dividends are given by \( Div_{t+1} = (R^L_{t+1}D_t - R^d_tD_t) \). The intermediary seeks to maximize its equity value which is given by:
\[ Q_t^i = E_t \sum_{j=1}^{\infty} \frac{\beta^j \Lambda_{t+j}}{\Lambda_t} \text{Div}_{t+j} \]  

where \( \Lambda_t \) is the marginal utility of real income for the representative household that owns the lender.

The first-order condition shows that in expectation, the lender makes zero profits, but ex post profits and losses can occur.\(^2\) We assume that losses are covered by households as negative dividends. This is similar to the standard assumption in the dynamic new-Keynesian (DNK) model, e.g., Woodford (2003). That is, the sticky price firms are owned by the household and pay out profits to the household. These profits are typically always positive (for small shocks) because of the steady state mark-up over marginal cost. Similarly, one could introduce a steady-state wedge (e.g., monopolistic competition among lenders) in the lender’s problem so that dividends are always positive. But this assumption would have no effect on the model’s dynamics so we dispense from it for simplicity.

**Entrepreneurs and the Loan Contract**

Entrepreneurs are the sole accumulators of physical capital. At the beginning of period \( t \), the entrepreneurs sell all of their accumulated capital to *capital agencies* at beginning-of-period capital price \( Q_{beg}^t \). At the end of the period, the entrepreneurs purchase the entire capital stock \( \bar{K}_t \), including any net additions to the stock, at end-of-period price \( Q_t \). This re-purchase of capital is financed with entrepreneurial net worth \( (NW_t) \) and external financing from a lender. The external finance takes the form of a one period loan contract. The gross return to holding capital from time-\( t \) to time \( t+1 \) is given by:

\(^2\) In contrast, BGG assume that bank profits are always zero ex post so that the lender’s return in pre-determined. This is not a feature of the optimal contract. See Carlstrom et al. (2013) for details.
Below we show that \( Q^\text{beg}_{t+1} = Q_{t+1} (1 - \delta) + \left[ \rho_{t+1} u_{t+1} - a(u_{t+1}) \right] \approx Q_{t+1} (1 - \delta) + \rho_{t+1} \), the latter term coinciding with the expression in BGG. Variations in \( R^k_{t+1} \) are the source of aggregate risk in the loan contract. The external financing is subject to a costly-state-verification (CSV) problem because of idiosyncratic risk. In particular, one unit of capital purchased at time-\( t \) is transformed into \( \omega_{t+1} \) units of capital in time \( t+1 \), where \( \omega_{t+1} \) is an idiosyncratic random variable with density \( \phi(\omega) \) and cumulative distribution \( \Phi(\omega) \).

The realization of \( \omega_{t+1} \) is directly observed by the entrepreneur, but the lender can observe the realization only if a monitoring cost is paid, a cost that is fraction \( \mu_{mc} \) of the project outcome. Assuming that the entrepreneur and lender are risk-neutral, Townsend (1979) demonstrates that the optimal contract between entrepreneur and intermediary is risky debt in which monitoring only occurs if the promised payoff is not forthcoming. Payoff does not occur for sufficiently low values of the idiosyncratic shock, \( \omega_{t+1} < \sigma_{t+1} \). Let \( R^p_{t+1} \) denote the promised gross rate-of-return so that \( R^p_{t+1} \) is defined by

\[
R^p_{t+1} (Q_{t+1} \bar{K}_t - NW_t) \equiv \sigma_{t+1} R^k_{t+1} Q_{t+1} \bar{K}_t.
\]

We find it convenient to express this in terms of the leverage ratio \( \bar{K}_t \equiv \frac{Q_{t+1} K_t}{NW_t} \) so that Equation 10 becomes

\[
R^p_{t+1} = \sigma_{t+1} R^k_{t+1} \frac{\bar{K}_t}{\bar{K}_t - 1}.
\]

With \( f(\sigma_{t+1}) \) and \( g(\sigma_{t+1}) \) denoting the entrepreneur’s share and lender’s share of the project outcome, respectively, the lender’s ex post realized \( t+1 \) return on the loan contract is defined as:

\[
R^k_{t+1} = \frac{R^k_{t+1} g(\sigma_{t+1}) Q_{t+1} \bar{K}_t}{Q_{t+1} \bar{K}_t - NW_t} = R^k_{t+1} g(\sigma_{t+1}) \frac{\bar{K}_t}{\bar{K}_t - 1}
\]
\[ f(\sigma) = \int_{\sigma} \omega \phi(\omega) d\omega - [1 - \Phi(\sigma)] \sigma \]  
\[ g(\sigma) = \left[1 - \Phi(\sigma)\right] \sigma + (1 - \mu_{m_\omega}) \int_{0}^{\sigma} \omega \phi(\omega) d\omega \]  

Recall that the lender’s stochastic discount factor comes from the household, and the lender’s return is linked to the return on deposits via 8:

\[ E_t R_{t+1}^L \Lambda_{t+1} = R_t^d E_t \Lambda_{t+1} \]  

As for the entrepreneur, Carlstrom, Fuerst and Paustian (2013) show that the entrepreneur’s value function is linear with a time-varying coefficient we denote by \( V_t \), where \( V_t \) satisfies:

\[ V_t = (1 - \gamma) + \beta \gamma \bar{K} E_t V_{t+1} R_{t+1}^k f(\sigma_{t+1}) \]  

Using this valuation and expression 15, the end-of-time-\( t \) contracting problem is thus given by:

\[ \max_{E_t V_{t+1} R_{t+1}^k \bar{K}} \left\{ E_t V_{t+1} R_{t+1}^k \bar{K}, f(\sigma_{t+1}) \right\} \]  

subject to

\[ E_t R_{t+1}^k \frac{\bar{K}}{\bar{K} - 1} \Lambda_{t+1} g(\sigma_{t+1}) \geq R_t^d E_t \Lambda_{t+1} \]  

An important observation is that the choice of \( \sigma_{t+1} \) can be made contingent on public information available in time \( t+1 \). Indexation of this type is optimal. After some rearrangement, the contract optimization conditions include:

\[ V_{t+1} f'(\sigma_{t+1}) = \left[ E_t V_{t+1} R_{t+1}^k f(\sigma_{t+1}) \right] \Lambda_{t+1} g'(\sigma_{t+1}) \]  

\[ (\bar{K} - 1) E_t V_{t+1} R_{t+1}^k f(\sigma_{t+1}) = \left[ E_t V_{t+1} R_{t+1}^k f(\sigma_{t+1}) \right] E_t \Lambda_{t+1} R_{t+1}^k g(\sigma_{t+1}) \]  

\[ E_t \Lambda_{t+1} R_{t+1}^k \frac{\bar{K}}{\bar{K} - 1} g(\sigma_{t+1}) = R_t^d E_t \Lambda_{t+1} \]
A key result is given by (19). The optimal monitoring cut-off \( \sigma_{t+1} \) is independent of \( \bar{K} \) innovations in \( R^K_{t+1} \). The second-order condition implies that the ratio \( \frac{-f'(\sigma_{t+1})}{g(\sigma_{t+1})} \) is increasing in \( \sigma_{t+1} \). Hence, another implication of the privately optimal contract is that \( \sigma_{t+1} \) (and thus the optimal repayment rate \( R^R_{t+1} \)) is an increasing function of (innovations in) the marginal utility of wealth \( \Lambda_{t+1} \), and a decreasing function of (innovations in) the entrepreneur’s valuation \( V_{t+1} \).

The monitoring cut-off implies the behavior of the repayment rate (see 11). In log deviations, Carlstrom, Fuerst, and Paustian (2013) show that the repayment rate under the optimal contract is given by:

\[
\hat{r}_t^p = r_{t-1}^d + \left(1 - \Theta_z \right) \left[ 1 - \nu (\kappa - 1) \right] \bar{k}_{t-1} + (\hat{r}_t^k - E_{t-1} \hat{r}_t^k) + \frac{1}{\Psi} (\hat{\lambda}_t - E_{t-1} \hat{\lambda}_t) - \frac{1}{\Psi} (\hat{v}_t - E_{t-1} \hat{v}_t) \tag{22}
\]

\[
\hat{v}_t = E_t \sum_{j=0}^{\infty} \beta^j \left( \Xi \hat{k}_{t+j} + \hat{r}_{t+j}^k \right) \tag{23}
\]

where the hatted lower case letters denote log deviations, and the positive constants \( \Xi \) and \( \Psi \) are defined in the Appendix. Innovations in \( \hat{v}_t \) will be driven by innovations in \( \hat{r}_t^k \). Hence, for parsimony we will estimate an indexation rule of the form:

\[
\hat{r}_t^p = \Omega_{t-1} + \chi_k (\hat{r}_t^k - E_{t-1} \hat{r}_t^k) + \chi_{\lambda} (\hat{\lambda}_t - E_{t-1} \hat{\lambda}_t) \tag{24}
\]

where \( \Omega_{t-1} \) are the pre-determined variables that affect the repayment rate, e.g., \( \bar{k}_{t-1} \). As noted in the example sketched in Section 2, different indexation values will have dramatic effects on the financial accelerator. The original BGG model assumes that the lender’s return was pre-determined. From Equations 11-12, this implies that \( \chi_{\lambda} = 0 \) and \( \chi_k \approx 0 \), where \( \chi_k \) is modestly negative (\( \chi_k = -0.01 \), in our benchmark calibration).

Entrepreneurs have linear preferences and discount the future at rate \( \beta \). Given the high return to internal funds, they will postpone consumption indefinitely. To limit net worth accumulation and ensure that there is a need for external finance in the long run, we assume that fraction \((1-\gamma)\) of the entrepreneurs die each period. Their
accumulated assets are sold and the proceeds transferred to households as consumption. Given the exogenous death rate, aggregate net worth accumulation is described by

\[ NW_t = gNW_{t-1} - R^k f(\sigma) \eta_{nw,t} \]  

(25)

where \( \eta_{nw,t} \) is an exogenous disturbance to the distribution of net worth. We assume it follows the stochastic process

\[ \log \eta_{nw,t} = \rho_{nw} \log \eta_{nw,t-1} + \epsilon_{nw,t}, \]  

(26)

where \( \epsilon_{nw,t} \) is i.i.d. \( N(0, \sigma^2_{nw}) \). Equation 25 implies that \( NW_t \) is determined by the realization of \( R^k \) and the response of \( \sigma \) to these realizations. \( NW_t \) then enters the contracting problem in time \( t \) so that the realization of \( R^k \) is propagated forward.

As in Christiano, Motto, and Rostagno (2010), and Gilchrist, Ortiz and Zakrajšek (2009), we also consider time variation in the variance of the idiosyncratic shock \( \omega_t \). The variance of \( \omega_t \) is denoted by \( \sigma_t \) and follows the exogenous stochastic process given by

\[ \log \sigma_t = \rho_{\sigma} \log \sigma_{t-1} + \epsilon_{\sigma,t}, \]  

(27)

Shocks to this variance will alter the risk premium in the model.

**FINAL GOOD PRODUCERS**

Perfectly competitive firms produce the final consumption good \( Y_t \) combining a continuum of intermediate goods according to the CES technology:

\[ Y_t = \left[ \int_0^1 Y_i(t)^{\lambda_{p,t}} di \right]^{1/(1+\lambda_{p,t})} \]  

(28)

The elasticity \( \lambda_{p,t} \) follows the exogenous stochastic process

\[ \log \lambda_{p,t} = (1 - \rho_p) \log \lambda_p + \rho_p \log \lambda_{p,t-1} + \epsilon_{p,t} - \theta_p \epsilon_{p,t-1}, \]  

(29)

where \( \epsilon_{p,t} \) is i.i.d. \( N(0, \sigma_p^2) \). Fluctuations in this elasticity are price markup shocks. Profit maximization and the zero profit condition imply that the price of the final good, \( P_t \), is the familiar CES aggregate of the prices of the intermediate goods.
INTERMEDIATE GOODS PRODUCERS

A monopolist produces the intermediate good \( i \) according to the production function

\[
Y_t(i) = \max \{ A_t^{-\alpha} K_t(i)^\alpha L_t(i)^{1-\alpha} - A_t Y_t^{1-\alpha} F; 0 \},
\]

where \( K_t(i) \) and \( L_t(i) \) denote the amounts of capital and labor employed by firm \( i \). \( F \) is a fixed cost of production, chosen so that profits are zero in steady state. The variable \( A_t \) is the exogenous non-stationary level of TFP progress. Its growth rate \( (z_t \equiv \Delta \ln A_t) \) is given by

\[
z_t = (1 - \rho_z) \gamma_z + \rho_z z_{t-1} + \varepsilon_{z,t},
\]

where \( \varepsilon_{z,t} \) is \( i.i.d. \) \( N(0, \sigma_z^2) \). The other non-stationary process \( \Upsilon_t \) is linked to the investment sector and is discussed below.

Every period a fraction \( \xi_p \) of intermediate firms cannot choose its price optimally, but resets it according to the indexation rule

\[
P_t(i) = P_{t-1}(i) \pi_t \pi^{-1-p},
\]

where \( \pi_t \equiv P_t/P_{t-1} \) is gross inflation and \( \pi \) is its steady state. The remaining fraction of firms chooses its price \( P_t(i) \) optimally, by maximizing the present discounted value of future profits

\[
E_t \left\{ \sum_{s=0}^{\infty} \xi_{p} \frac{\beta^s \Lambda_{t+s} / P_{t+s}}{\Lambda_t / P_t} \left[ P_t(i) \left( \prod_{k=1}^{s} \pi_{k+t-1} \pi^{-1-t} \right) Y_{t+s}(i) - W_{t+s} L_{t+s}(i) - P_{t+s} \rho_{t+s} K_{t+s}(i) \right] \right\}
\]

where the demand function comes from the final goods producers, \( \Lambda_t / P_t \) is the marginal utility of nominal income for the representative household, and \( W_t \) is the nominal wage.

EMPLOYMENT AGENCIES

Firms are owned by a continuum of households, indexed by \( j \in [0,1] \). Each household is a monopolistic supplier of specialized labor, \( L_t(j) \), as in Erceg et al. (2000). A large number of competitive employment agencies combine this specialized labor into a homogenous labor input sold to intermediate firms, according to
As in the case of the final good, the desired markup of wages over the household’s marginal rate of substitution, \( \lambda_{w,t} \), follows the exogenous stochastic process

\[
\log \lambda_{w,t} = (1 - \rho_w) \log \lambda_{w} + \rho_w \log \lambda_{w,t-1} + \epsilon_{w,t} - \theta_w \epsilon_{w,t-1},
\]

where \( \epsilon_{w,t} \) is i.i.d. \( N(0, \sigma^2_w) \). This is the wage markup shock. Profit maximization by the perfectly competitive employment agencies implies that the wage paid by intermediate firms for their homogenous labor input is

\[
W_t = \left[ \int_0^1 W_t(j)^{-1/\lambda_{w,t}} \, dj \right]^{-\lambda_{w,t}}
\]

**CAPITAL AGENCIES**

The capital stock is managed by a collection of perfectly competitive capital agencies. These firms are owned by households and discount cash flows with \( \Lambda_t \), the marginal utility of real income for the representative household. At the beginning of period \( t \), these agencies purchase the capital stock \( K_{t-1} \) from the entrepreneurs at beginning-of-period price \( Q_{t}^{beg} \). The agencies produce capital services by varying the utilization rate \( u_t \) which transforms physical capital into effective capital according to

\[
K_t = u_t \bar{K}_{t-1}.
\]

Effective capital is then rented to firms at the real rental rate \( \rho_t \). The cost of capital utilization is \( a(u_t) \) per unit of physical capital. The capital agency then re-sells the capital to entrepreneurs at the end of the period at price \( Q_t \). The profit flow is thus given by:

\[
Q_t(1 - \delta) \bar{K}_{t-1} + [\rho_t u_t - a(u_t)] \bar{K}_{t-1} - Q_t^{beg} \bar{K}_{t-1}
\]

Profit maximization implies

\[
Q_t^{beg} = Q_t(1 - \delta) + [\rho_t u_t - a(u_t)]
\]

\[
\rho_t = a'(u_t)
\]
In steady state, \( u = 1, a(1) = 0 \) and \( \delta = a''(1)/a'(1) \). Hence, in the neighbourhood of the steady state

\[
Q^\text{log} \approx Q(1-\delta) + \rho_t
\]

which is consistent with BGG's definition of the intertemporal return to holding capital

\[
R^t_i \equiv \frac{\rho_t + (1-\delta)Q_t}{Q_{t-1}}.
\]

**NEW CAPITAL PRODUCERS**

New capital is produced according to the production technology that takes \( I_t \) investment goods and transforms them into \( \mu_t \left[ 1 - S \left( \frac{I_t}{I_{t+1}} \right) \right] I_t \) new capital goods. The time-\( t \) profit flow is thus given by

\[
Q_t \mu_t \left[ 1 - S \left( \frac{I_t}{I_{t+1}} \right) \right] I_t - P_t^t I_t
\]

where \( P_t^t \) is the relative price of the investment good. The function \( S \) captures the presence of adjustment costs in investment, as in Christiano et al. (2005). The function has the following convenient steady state properties: \( S = S' = 0 \) and \( S'' > 0 \). These firms are owned by households and discount future cash flows with \( \Lambda_t \), the marginal utility of real income for the representative household. JPT refer to the investment shock \( \mu_t \) as a shock to the marginal efficiency of investment (MEI) as it alters the transformation between investment and installed capital. JPT conclude that this shock is the primary driver of output and investment at business cycle frequencies. The investment shock follows the stochastic process

\[
\log \mu_t = \rho_{\mu} \log \mu_{t-1} + \epsilon_{\mu,t},
\]

where \( \epsilon_{\mu,t} \) is i.i.d. \( N \left( 0, \sigma_{\mu}^2 \right) \).
INVESTMENT PRODUCERS

A competitive sector of firms produces investment goods using a linear technology that transforms one consumption good into $Y_t$ investment goods. The exogenous level of productivity $Y_t$ is non-stationary with a growth rate ($\nu_t = \Delta \log Y_t$) given by

$$\nu_t = (1 - \rho) \nu_{t-1} + \rho \nu_{t-1} + \epsilon_{t-1}.$$  \hspace{1cm} (44)

The constant returns production function implies that the price of investment goods (in consumption units) is equal to $\frac{1}{Y_t}$.

HOUSEHOLDS

Each household maximizes the utility function

$$E_t \left\{ \sum_{s=0}^{\infty} \beta^s b_{t+s} \left[ \ln (C_{t+s} - hC_{t+s-1}) - \varphi \frac{L_{t+s}(j)^{1+\psi}}{1+\psi} \right] \right\},$$  \hspace{1cm} (45)

where $C_t$ is consumption, $h$ is the degree of habit formation and $b_t$ is a shock to the discount factor. This intertemporal preference shock follows the stochastic process

$$\log b_t = \rho b_{t-1} + \epsilon_{b,t},$$  \hspace{1cm} (46)

where $\epsilon_{b,t}$ is i.i.d. $N(0, \sigma_b^2)$. Since technological progress is nonstationary, utility is logarithmic to ensure the existence of a balanced growth path. The existence of state contingent securities ensures that household consumption is the same across all households. The household’s flow budget constraint is

$$C_t + T_t + D_t + \frac{B_t}{P_t} - \frac{R_{t-1} B_{t-1}}{P_t} + \frac{W_t(j)}{P_t} L_t(j) + D_{t-1} R_{t-1} + \text{profits}_t,$$  \hspace{1cm} (47)

where $D_t$ denotes real deposit at the lender, $T_t$ is lump-sum taxes, and $B_t$ is holdings of nominal government bonds that pay gross nominal rate $R_t$. The term $\text{profits}_t$ denotes the combined profit flow of all the firms owned by the representative agent including lenders, intermediate goods producers, capital agencies, and new capital producers. Every period a fraction $\xi$ of households cannot freely set its wage, but follows the indexation rule
\[ W_t(j) = W_{t-1}(j)(\pi \frac{\pi_t}{\pi_t + \alpha (\frac{z_t}{1-\alpha})^{\frac{1}{\alpha}}} \gamma_t (\pi \frac{\pi_t}{\pi_t + \alpha (\frac{z_t}{1-\alpha})^{\frac{1}{\alpha}}})^{1-\gamma_t}, \] (48)

The remaining fraction of households chooses instead an optimal wage \( W_t(j) \) by maximizing

\[ E_1 \left\{ \sum_{i=0}^{\infty} \xi^i \beta^i \left[ -b_{i+1} \varphi \frac{L_{i+1}(j)^{1+\psi}}{1+\psi} + \Lambda_{i+1} W_t(j) L_{i+1}(j) \right] \right\} \] (49)

subject to the labor demand function coming from the firm.

**THE GOVERNMENT**

A monetary policy authority sets the nominal interest rate following a feedback rule of the form

\[ \frac{R_t}{R} = \left( \frac{R_{t-1}}{R} \right)^{\rho R} \left[ \left( \frac{\pi_t}{\pi_t} \right)^{\phi R} \left( \frac{X_t}{X_t^*} \right)^{\phi L} \left( \frac{X_t}{X_t^*} / X_t^{*1-\gamma_t} \right)^{\gamma_t} \right] \eta_{mp,t}, \] (50)

where \( R \) is the steady state of the gross nominal interest rate. The interest rates respond to deviations of inflation from its steady state, as well as to the level and the growth rate of the GDP gap \( (X_t / X_t^*) \). The monetary policy rule is also perturbed by a monetary policy shock, \( \eta_{mp,t} \), which evolves according to

\[ \log \eta_{mp,t} = \rho_{mp} \log \eta_{mp,t-1} + \varepsilon_{mp,t}, \] (51)

where \( \varepsilon_{mp,t} \) is i.i.d. \( N(0,\sigma_{mp}^2) \). Public spending is determined exogenously as a time-varying fraction of output.

\[ G_t = \left( 1 - \frac{1}{g_t} \right) Y_t, \] (52)

where the government spending shock \( g_t \) follows the stochastic process

\[ \log g_t = (1 - \rho_g) \log g + \rho_g \log g_{t-1} + \varepsilon_{g,t}, \] (53)

with \( \varepsilon_{g,t} \sim i.i.d. N\left(0,\sigma_g^2\right) \). The spending is financed with lump sum taxes.

**MARKET CLEARING**

The aggregate resource constraints are given by:
\[ C_t + \frac{I_t}{Y_t} + G_t + a(u_t)_{t-1} = Y_t \]  

\[ \bar{K}_t = (1-\delta) \left( 1 - \mu_{mc} \frac{\gamma}{\omega} \int_0^\infty f(\omega) d\omega \right) \bar{K}_{t-1} + \mu \left[ 1 - S\left( \frac{I_t}{I_{t-1}} \right) \right] I_t. \]

This completes the description of the model. We now turn to the estimation of the linearized model.

4. ESTIMATION

The linearized version of the model equations are collected in the appendix. The three fundamental agency cost parameters are the steady state idiosyncratic variance (\(\sigma_{ss}\)), the entrepreneurial survival rate (\(\gamma\)), and the monitoring cost fraction (\(\mu_{mc}\)). In contrast to DeGraeve (2008), we follow Christiano et al. (2010), and calibrate these parameters to be consistent with long run aspects of US financial data. We follow this calibration approach because these parameters are pinned down by long run or steady state properties of the model, not the business cycle dynamics that the Bayesian estimation is trying to match. In any event, these three parameters are calibrated to match the steady state levels of the risk premium (\(R^p - R^d\)), leverage ratio (\(\bar{K}\)), and default rate (\(\Phi(\sigma_{ss})\)). In particular, they are chosen to deliver a 200 bp annual risk premium (BAA-Treasury spread), a leverage ratio of \(\bar{K} = 1.95\), and a quarterly default rate of 0.03/4. These imply an entrepreneurial survival rate of \(\gamma = 0.98\), a standard deviation of \(\sigma_{ss} = 0.28\), and a monitoring cost of \(\mu_{mc} = 0.12\). A key expression in the log-linearized model is the reduced-form relationship between the risk premium and leverage:

\[ E_t \hat{r}_{t+1}^k - \hat{r}_t^d = v \left( \hat{q}_t + \hat{k}_t - \hat{n}_t \right) + \hat{\sigma}_t \]
(See the Appendix for details.) The value of $\nu$ implied by the previous calibration is $\nu = 0.041$. This is thus imposed in the estimation of the financial models. For JPT we have $\nu = 0$. Steady state relationships also imply that we calibrate $\delta = 0.025$, and $(1-1/g) = 0.22$. The remaining parameters are estimated using familiar Bayesian techniques as in JPT. For the non-financial parameters of the model we use the same priors as in JPT.

We treat as observables the growth rates of real GDP, consumption, investment, the real wage, and the relative price of investment. The other observables include employment, inflation, the nominal rate, leverage, and the risk premium. Employment is measured as the log of per capita hours. Inflation is the consumption deflator, and the nominal rate is the federal funds rate. The series for leverage comes from Gilchrist, Ortiz and Zakrajsek (2009). The risk premium is the spread between the BAA and ten year Treasury. The time period for the estimation is 1954:3-2009:1. We choose the end of the sample period to avoid the observed zero bound on the nominal rate.

We estimate four versions of the model. Along with all the exogenous shocks outlined in the paper, we also include autocorrelated measurement error between the model’s risk premium and the observed risk premium. Autocorrelated measurement error is also included for leverage. The first model we label JPT as it corresponds to the model without agency costs ($\nu = 0$). Note that to match the observed financial variables, the JPT model will assign all risk premium and leverage variation to autocorrelated measurement error. The remaining three models have operative agency costs ($\nu = 0.041$). Recall that the optimal contract has the form given in Equation 24. Our three estimates consider variations on this basic form. In the model labeled BGG we impose the level of indexation implicitly assumed by BGG: $\chi_k = -0.01$, $\chi_\lambda = 0$. For the model labeled $R^k$-indexation, we set $\chi_\lambda = 0$ and estimate the value of $\chi_k$. For the model labeled $R^k$ & $\lambda$-indexation, we estimate both indexation parameters. We use diffuse priors on the

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3 As a form of sensitivity analysis, we also estimated in the financial models. We found that the estimation is quite sensitive to priors, again suggesting that it is not well identified by business cycle dynamics.
indexation parameters with a uniform distribution centered at 0 and with a standard deviation of 2.

The agency cost models also include two financial shocks: 

- time-varying movements in idiosyncratic risk, and
- exogenous redistributions of net worth. Both of these shocks are irrelevant in the JPT model in which lending is not subject to the CSV problem. We posit priors for the standard deviation and autocorrelation of these financial shocks in a manner symmetric with the non-financial exogenous processes in JPT.

The estimation results are summarized in Table 1. The BGG, $R^k$-indexation, and $R^k$ & $\lambda$-indexation agency cost models dominate the JPT model as the JPT model cannot capture the forecastability of leverage and the risk premium that is in the data. Comparing BGG and $R^k$-indexation, the data rejects the BGG level of indexation preferring a level of contract indexation that is economically significant: $\chi_k = 1.70$ with a 90% confidence interval between 1.36 and 2.05. Results are on a similar range in the case of the $R^k$ & $\lambda$-indexation estimation with $\chi_k = 2.23$ with a 90% confidence interval between 1.52 and 2.99. As suggested by the example in Section 2, this level of indexation will imply significantly different responses to shocks compared to the BGG assumption. We will see this manifested in the IRF below. The estimated level of indexation to the marginal utility of wealth is $\chi_\lambda = 0.67$ with a 90% confidence interval between 0 and 1.38. The combination of the two indexation parameters under the $R^k$ & $\lambda$-indexation specification generates dynamics that are similar to those of the $R^k$-indexation specification alone.

Two other differences in parameter estimates are worth some comment. First, the BGG model estimates a significantly smaller size for investment adjustment costs (\(S^n\)) in the table: \(S^n = 1.68\) for BGG, but 2.99 for $R^k$-indexation, 2.50 for $R^k$ & $\lambda$-indexation, and 2.92 for JPT. The level of adjustment costs has two contrasting effects. First, lower adjustment costs will increase the response of investment to aggregate shocks. Second, lower adjustment costs imply smaller movements in the price of installed capital (\(Q_t\)) and thus smaller financial accelerator effects in the BGG model.
A second important difference in parameter estimates is in the standard deviation of the shocks. Compared to JPT, the BGG model estimates a significantly smaller volatility in the MEI shocks, and instead shifts this variance on to net worth shocks. Recall that the principle conclusion of JPT is the importance of the MEI shocks in the business cycle. But we once again end up with the JPT conclusion with regards to the importance of MEI shocks in the $R^k$-indexation and $R^k$ & $\lambda$-indexation models. An interesting question we take up below is why the BGG model downplays these shocks so significantly.

Table 2 reports the variance decomposition of three key variables: GDP, investment, and the risk premium. The JPT results are replicated here: the MEI shocks account for a substantial amount of business cycle variability in GDP (60% at the 8-quarter horizon) and investment (77% at the 8-quarter horizon). This conclusion is largely unchanged with $R^k$-indexation and $R^k$ & $\lambda$-indexation. Evidently the estimated level of indexation results in real behaviour similar to a model without agency costs. This is particularly clear in the IRFs presented in Figure 3 that we discuss below.

In contrast to the $R^k$-indexation and $R^k$ & $\lambda$-indexation models, BGG places much less weight on the MEI shocks and instead shifts this variance to the financial shocks (the idiosyncratic variance and net worth shocks) and the monetary policy shock. For the case of investment at the 8-quarter horizon, the BGG model places 15% of the variance on the MEI shocks (compared to 68% for the $R^k$-indexation and $R^k$ & $\lambda$-indexation models, and 77% for JPT). The importance of the two financial shocks increases from 13% under $R^k$ & $\lambda$-indexation and 14% under $R^k$-indexation, to 44% for BGG. The estimated level of financial shocks depends critically upon the estimated level of indexation.

The advantage of the financial models is showcased in the variance decomposition of the risk premium. By assumption, JPT assigns 100% of this variation to measurement error. In contrast, the financial models explain large portions of the risk premia movement by forces within the model. For example, at the 8-quarter horizon, the $R^k$-indexation model assigns less than 50% to measurement error, and the $R^k$ & $\lambda$-indexation assigns less than 10% to measurement error. This predictability of the risk premium is echoed by De Graeve (2008).
Why does the BGG model downplay the MEI shocks and thus shift variance to the other shocks? The answer is quite apparent from Figure 3a. The figure sets all parameter values to those estimated in the $R^k$-indexation model, except for the levels of indexation ($\lambda_k \approx 0$ for BGG and $\lambda_k = 2.23$ and $\lambda_k = 0.67$ for the $R^k$ & $\lambda$-indexation model), and the level of agency cost effects ($\nu = 0$ for JPT). A positive innovation in MEI leads to a fall in the price of capital. Since the BGG contract is not indexed to the return to capital, the shock leads to a sharp decline in entrepreneurial net worth, and thus a sharp increase in the risk premium. This procyclical movement in the risk premium is in sharp contrast to the data. Hence, the Bayesian estimation in the BGG model estimates only a small amount of variability coming from the MEI shocks. Notice that in the $R^k$-indexation and $R^k$ & $\lambda$-indexation models net worth is almost unchanged in response to an MEI shock, so that the impact effect on the risk premium is countercyclical. The main difference among models is the behavior of the repayment that under credit contract indexation is lowered in response to the drop in the return on capital, while in BGG it remains unchanged. The $R^k$-indexation and $R^k$ & $\lambda$-indexation models are thus consistent with MEI shocks driving the cycle, and the risk premium being countercyclical. The similarity of the $R^k$-indexation and JPT model is also apparent: the two IRFs to an MEI shock largely lie on top of one another.

Since the BGG model downplays the importance of MEI shocks, and shifts this variance to other shocks. Figures 3b-3c plot the IRFs to the two financial shocks. The good news with the two financial shocks is that the spread is now countercyclical. But the difficulty with the financial shocks is that they result in countercyclical consumption. This is the familiar co-movement puzzle that arises when a positive shock in one sector (e.g., higher net worth mitigates agency costs in capital accumulation) leads to a downward production movement in the other sector. However, this comovement problem does not arise with risk premium shocks in the $R^k$-indexation and $R^k$ & $\lambda$-indexation models. As an

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4 Alternatively we could have considered the IRFs for each model at each model’s parameter estimates. These IRFs are similar to those reported here, but we find Figure 5 more intuitive as it is holding all other parameters fixed except for the degree of indexation and the presence of agency costs.
aside, note that a shock to net worth has a larger effect on net worth and capital prices in the BGG model. This is just a manifestation of the multiplier intuition outlined in Section 2.

Figure 3d plots the IRF to a monetary shock. In the case of BGG, the IRFs exhibit plausible comovement and countercyclical spreads. The BGG estimation does not put more weight on these policy shocks because the funds rate is an observable, and thus limits possible interest rate variability. In contrast, it is quite clear why the Indexation model puts so little weight on monetary policy shocks. In the case of Indexation, the spread is procyclical, a clear counterfactual prediction.

As a form of sensitivity analysis, Table 3 presents the estimation results for \textit{i.i.d.} measurement error in the financial variables. The $R^k$ & $\lambda$-indexation now wins the model horse race, with JPT coming in significantly worse than the three financial models. The degree of $R^k$-indexation is much larger than the case with autocorrelated measurement error. Further the level of indexation to the marginal utility of wealth is estimated to be significantly positive, in line with the theory outlined above.

5. CONCLUSION

This paper began as an empirical investigation of the importance of agency costs and contract indexation in the business cycle. To reiterate, our principle results include the following. First, the financial models appear to be an improvement over the financial-frictionless JPT. Second, $R^k$-indexation appears to be an important characteristic of the data. Third, the importance of financial shocks (net worth and idiosyncratic variance) in explaining the business cycle is significantly affected by the estimated degree of $R^k$-indexation. In short, we find evidence for the importance of financial shocks in the business cycle. But the evidence also suggests that the effect of non-financial shocks on real activity is unaffected by the inclusion of financial forces in the model. That is, the results suggest the importance of financial shocks, but not the existence of a financial accelerator. This analysis thus implies that Bayesian estimation of
financial models should include estimates of contract indexation. Empirical analyses that impose zero contract indexation likely distort both the source of business cycle shocks and their transmission mechanism.
APPENDIX

1. Linearized System of Equations:

\[
\hat{y}_t = \frac{y^+ F}{y} \left[ \alpha \hat{k}_t + (1 - \alpha) \hat{L}_t \right] \tag{A1}
\]

\[
\hat{\rho}_t = \hat{w}_t + \hat{L}_t - \hat{k}_t \tag{A2}
\]

\[
\hat{s}_t = \alpha \hat{\rho}_t + (1 - \alpha) \hat{w}_t \tag{A3}
\]

\[
\hat{\pi}_t = \frac{\beta}{1 + \beta t_p} E_t \hat{\pi}_{t+1} + \frac{t_p}{1 + \beta t_p} \hat{\pi}_{t-1} + \frac{(1 - \beta \xi_p) (1 - \xi_p)}{(1 + \beta t_p) \xi_p} \hat{s}_t + \hat{\lambda}_{p,t} \tag{A4}
\]

\[
\hat{\lambda}_t = \frac{h \beta e^{\tau_z}}{(e^{\tau_z} - h)} E_t \hat{\lambda}_{t+1} - \frac{e^{\tau_z} + h \beta}{(e^{\tau_z} - h)} \hat{c}_t + \frac{h e^{\tau_z}}{(e^{\tau_z} - h)} \hat{c}_{t-1} + \frac{h \beta e^{\tau_z} \rho_z - h e^{\tau_z}}{(e^{\tau_z} - h)} \hat{z}_t + \frac{e^{\tau_z} - h \beta \rho_z}{(e^{\tau_z} - h)} \hat{b}_t + \left[ \frac{h \beta e^{\tau_z} \rho_z - h e^{\tau_z}}{(e^{\tau_z} - h)} \right] \left( \frac{\alpha}{1 - \alpha} \right) \hat{u}_t \tag{A5}
\]

\[
\hat{\lambda}_t = \hat{R}_t + E_t (\hat{\lambda}_{t+1} - \hat{\pi}_{t+1} - \hat{\pi}_{t-1} - \frac{\alpha}{1 - \alpha} \hat{v}_{t+1}) \tag{A6}
\]

\[
\hat{\rho}_t = \theta \hat{u}_t \tag{A7}
\]

\[
E_t \hat{\pi}_{t+1} = \hat{\lambda}_t - E_t \hat{\lambda}_{t+1} + E_t \hat{\pi}_{t+1} + \left( \frac{\alpha}{1 - \alpha} \right) E_t \hat{v}_{t+1} \tag{A8}
\]

\[
\hat{q}_t = -\hat{\mu}_t + e^{2(\gamma_z + \gamma_s)} S \left( \hat{i}_t - \hat{i}_{t-1} + \hat{z}_t + \frac{1}{1 - \alpha} \hat{u}_t \right) \tag{A9}
\]

\[
-\beta e^{2(\gamma_z + \gamma_s)} S E_t \left( \hat{i}_{t+1} - \hat{i}_t + \hat{z}_{t+1} + \frac{1}{1 - \alpha} \hat{u}_{t+1} \right) \]

\[
\hat{k}_t = \hat{u}_t + \hat{k}_{t-1} - \hat{z}_t - \frac{1}{1 - \alpha} \hat{u}_t \tag{A10}
\]

\[
\hat{k}_t = (1 - \delta) e^{-\gamma_z} \left( \hat{k}_{t-1} - \hat{z}_t - \frac{1}{1 - \alpha} \hat{u}_t \right) + \left[ 1 - (1 - \delta) e^{-\gamma_z} \right] (\hat{\mu}_t + \hat{i}_t) \tag{A11}
\]

\[
\hat{w}_t = \frac{1}{1 + \beta} \hat{w}_{t-1} + \frac{\beta}{1 + \beta} E_t \hat{w}_{t+1} - \frac{(1 - \beta \xi_w) (1 - \xi_w)}{(1 + \beta t_w) \xi_w} \hat{g}_{w,t} + \frac{t_w}{1 + \beta} \hat{\pi}_{t-1} \tag{A12}
\]
estimating contract indexation in a financial accelerator model

\[ -\frac{1+\beta t_{w}}{1+\beta} \hat{\pi}_{t} + \frac{\beta}{1+\beta} E_{t} \hat{\pi}_{t+1} + \frac{t_{w}}{1+\beta} \left( \hat{z}_{t-1} + \left( \frac{\alpha}{1-\alpha} \right) \hat{u}_{t-1} \right) - \frac{1+\beta t_{w} - \beta \rho_{z}}{1+\beta} \hat{\pi}_{t} + \hat{\lambda}_{w,t} \]  

(A12)

\[ \hat{g}_{w,t} = \hat{w}_{t} - (\psi \hat{L}_{t} + \hat{b}_{t} - \hat{\lambda}_{t}) \]  

(A13)

\[ \hat{R}_{t} = \rho_{x} \hat{R}_{t-1} + (1 - \rho_{x}) \left[ \phi_{x} \hat{x}_{t} + \phi_{x} \left( \hat{x}_{t} - \hat{x}_{t}^{*} \right) \right] + \phi_{x} \left[ \left( \hat{x}_{t} - \hat{x}_{t-1} \right) - \left( \hat{x}_{t}^{*} - \hat{x}_{t-1}^{*} \right) \right] + \hat{n}_{mx,t} \]  

(A14)

\[ \hat{x}_{t} = \hat{y}_{t} - \frac{\rho k}{y} \hat{u}_{t} \]  

(A15)

\[ \frac{1}{g} \hat{y}_{t} = \frac{1}{g} \hat{g}_{t} + \frac{c}{y} \hat{c}_{t} + \frac{i}{y} \hat{i}_{t} + \frac{\rho k}{y} \hat{u}_{t} \]  

(A16)

\[ \hat{r}_{d}^{i} = \hat{R}_{t} - E_{t} \hat{\pi}_{t+1} \]  

(A17)

\[ \hat{r}_{i}^{k} = \beta e^{-(\gamma_{x} + \gamma_{y})} (1 - \delta) \hat{\pi}_{i} + \left[ 1 - \beta e^{-(\gamma_{x} + \gamma_{y})} (1 - \delta) \right] \hat{\lambda}_{i} - \hat{q}_{t-1} \]  

(A18)

For the agency cost model, we replace (A8) with

\[ E_{t} \hat{r}_{t+1}^{k} = \hat{\lambda}_{t} - E_{t} \hat{\pi}_{t+1} + E_{t} \hat{z}_{t+1} + \left( \frac{\alpha}{1-\alpha} \right) E_{t} \hat{u}_{t} + E_{t} \hat{\pi}_{t} + \theta_{g} \left( \hat{q}_{t} + \hat{\pi}_{t+1} + \hat{\lambda}_{t} - \hat{n}_{t+1} \right) + \hat{\sigma}_{t} \]  

(A8’)

And add the following equations:

\[ \hat{n}_{t} = \frac{\gamma_{y}}{\beta} \left( \hat{r}_{i}^{k} - \hat{r}_{i}^{l} \right) + \frac{\gamma_{y}}{\beta} \left( \hat{r}_{i}^{l} + \hat{n}_{t-1} \right) + \gamma_{x} \frac{r P}{\beta} \left( \hat{k}_{t} + \hat{q}_{t} + \hat{r}_{i}^{k} \right) - \hat{z}_{t} - \left( \frac{1}{1-\alpha} \right) \hat{u}_{t} + \hat{n}_{mx,t} \]  

(A19)

\[ \hat{r}_{i}^{l} = \hat{r}_{i-1}^{l} + \left[ 1 + \theta_{g} \left( \chi_{k} - 1 \right) \right] \left( \hat{r}_{i}^{k} - E_{t} \hat{r}_{t+1}^{k} \right) + \theta_{g} \alpha \left( \hat{\lambda}_{t} - E_{t} \hat{\lambda}_{t} \right) \]  

(A20)

2. The Derivation of A8’ and A20

The optimal contract (19)-(21) can be expressed as

\[ V_{t+1, f} (\omega_{t+1}) = \left[ \frac{E_{t+1, f} (\omega_{t+1})}{E_{t+1, g} (\omega_{t+1})} \right] \Lambda_{t+1, g} (\omega_{t+1}) \]  

(A21)

\[ \bar{K}_{t} E_{t} R_{t+1}^{k} V_{t+1, f} (\omega_{t+1}) = \frac{-E_{t+1, f} (\omega_{t+1})}{E_{t+1, g} (\omega_{t+1})} R_{t}^{d} E_{t} \Lambda_{t+1} \]  

(A22)

\[ E_{t} \Lambda_{t+1, R_{t+1}^{k} \bar{K}_{t}} \Lambda_{t+1, g} (\omega_{t+1}) = R_{t}^{d} E_{t} \Lambda_{t+1} \]  

(A23)
It is convenient to define $F(\sigma_{t+1}) = -\frac{f'(\sigma_{t+1})}{g'(\sigma_{t+1})}$, where $\Psi \equiv \frac{\sigma_{ss} f'(\sigma_{ss})}{F(\sigma_{ss})} > 0$, by the second order condition. Linearizing (A21)-(A23) we have

\begin{align*}
\Psi(\sigma_{t+1} - E_1 \sigma_{t+1}) &= (\lambda_{t+1} - E_1 \lambda_{t+1}) + (\nu_{t+1} - E_1 \nu_{t+1}) \\
E_t(\hat{r}^k_{t+1} - \hat{r}^d_{t+1}) + \kappa_t &= (\Psi - \theta_f) E_1 \sigma_{t+1} \\
E_t(\hat{r}^k_{t+1} - \hat{r}^d_{t+1}) &= \left(\frac{1}{\kappa - 1}\right) \kappa_t - \theta_g E_1 \sigma_{t+1}
\end{align*}

where $\theta_g \equiv \frac{\sigma_{ss} g'(\sigma_{ss})}{g(\sigma_{ss})}$, with $0 < \theta_g < 1$, and $\theta_f \equiv \frac{\sigma_{ss} f'(\sigma_{ss})}{f(\sigma_{ss})} < 0$. Solving (A25)-(A26) we have:

\begin{align*}
E_t \sigma_{t+1} &= \frac{\kappa}{\kappa - 1} \left(\Psi - \theta_f + \theta_g\right) \kappa_t \\
E_t(\hat{r}^k_{t+1} - \hat{r}^d_{t+1}) &= \left[\frac{(\Psi - \theta_f + \theta_g) - \kappa \theta_g}{(\kappa - 1)(\Psi - \theta_f + \theta_g)}\right] \kappa_t \equiv \nu \kappa_t
\end{align*}

Using the definition of leverage and the deposit rate, (A28) is the same as (A8'). The linearized lender return and promised payment are given by:

\begin{align*}
\hat{r}^l_{t+1} &= \hat{r}^k_{t+1} + \theta_g \sigma_{t+1} - \left(\frac{1}{\kappa - 1}\right) \kappa_t \\
\hat{r}^p_{t+1} &= \hat{r}^k_{t+1} + \sigma_{t+1} - \left(\frac{1}{\kappa - 1}\right) \kappa_t
\end{align*}

Combining (A24) and (A30) we have:

\begin{align*}
\hat{r}^p_{t+1} &= \hat{r}^d_{t+1} + \frac{1 - \Theta_g}{\Theta_g (\kappa - 1)} \left[1 - \nu (\kappa - 1)\right] \kappa_{t-1} + (\hat{r}^k_{t+1} - E_t \hat{r}^d_{t+1}) + \frac{1}{\Psi} (\hat{\lambda}_t - E_{t-1} \hat{\lambda}_t) - \frac{1}{\Psi} (\hat{\nu}_t - E_{t-1} \hat{\nu}_t)
\end{align*}

where from (16) we have
\[
\hat{v}_t = E_t \sum_{j=0}^{\infty} \beta^{j+1} \left( \Xi \hat{k}_{t+j} + \hat{r}_{t+j+1}^k \right) \tag{A32}
\]

where \( \Xi \equiv \left[ 1 + \frac{\beta}{\gamma} \left( v - \frac{1}{(\kappa-1)} \right) \right] \). Innovations in \( \hat{v}_t \) are dominated by innovations in \( \hat{r}_t^k \), so for parsimony we will estimate a promised payment of the form

\[
\hat{r}_t^p = E_{t-1} \hat{r}_t^p + \chi(\hat{r}_t^k - E_{t-1} \hat{r}_t^k) + \chi(\lambda_t - E_{t-1} \lambda_t) \tag{A33}
\]

Combining this with (A29)-(A30) we have:

\[
\hat{r}_t^l = \hat{r}_{t-1}^d + [1 + \theta_g (\chi_k - 1)](\hat{r}_t^k - E_{t-1} \hat{r}_t^k) + \theta_g \chi(\lambda_t - E_{t-1} \lambda_t) \tag{A34}
\]

This is just (A20).
Figure 1
The Multiplier as a Function of Indexation.

Figure 2
A Shock to Asset Demand

(Asset price is blue line. Net worth evolution is red line.) Demand shock shifts up asset price. The new equilibrium in \((n,q)\) space depends upon the level of indexation. Lower levels of indexation amplify these effects.
Figure 3.a. Impulse Response Functions to a One Standard Deviation Marginal Efficiency of Investment Shock

Keeping parameters constant to the $R^2$ indexation model except $R^2$ and lambda indexation parameters ($\chi_k$ and $\chi_\lambda$)
Figure 3.b. Impulse Response Functions to a One Standard Deviation Net Worth Shock

Keeping parameters constant to the Rk indexation model except Rk and lambda indexation parameters (χk and χλ).

- Output
- Consumption
- Investment
- Inflation
- Wages
- Federal Funds
- Spread
- Net Worth
- Price of Capital
- Return on Capital (Rk)
- Promised Repayment
- Marginal Utility of Wealth (λ)

BGG

Indexation to Rk

Indexation to Rk and λ
Figure 3.c: Impulse Response Functions to a One Standard Deviation Idiosyncratic Variance Shock

Keeping parameters constant to the $R^k$ indexation model except $R^k$ and lambda indexation parameters ($\chi_k$ and $\chi_\lambda$)

---

Output

---

Consumption

---

Investment

---

Inflation

---

Wages

---

Federal Funds

---

Spread

---

Net Worth

---

Price of Capital

---

Return on Capital ($R^k$)

---

Promised Repayment

---

Marginal Utility of Wealth ($\lambda$)

---

BGG

---

Indexation to $R^k$

---

Indexation to $R^k$ and $\lambda$
Figure 3.d. Impulse Response Functions to a One Standard Deviation Monetary Policy Shock
Keeping parameters constant to the R^k indexation model except R^k and lambda indexation parameters (q and \chi).

Output
Consumption
Investment
Inflation
Wages
Federal Funds
Spread
Net Worth
Price of Capital
Return on Capital (R^k)
Promised Repayment
Marginal Utility of Wealth (\lambda)

_______ JPT
--------- BGG
Indexation to R^k
Indexation to R^k and \lambda
Table 1: Models Estimations and Models Comparisons with BAA - T10 credit spread and leverage data.

All models have subordinated measurement errors in the credit spread and leverage series.

<table>
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<tr>
<th>Coefficient</th>
<th>Description</th>
<th>Prior</th>
<th>Posterior 1</th>
<th>Posterior 2</th>
<th>Posterior 3</th>
<th>Posterior 4</th>
<th>Standard deviation</th>
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<td>0.15</td>
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Note: Differenced coefficients, *p* ≤ 1.50 are set to be 0.05 or 0.10. 

For the agency cost models (BGG and Indexation) the following parameters are also calibrated: entrepreneurial survival rate, e, and the indexation to the marginal utility of wealth parameter, λ, is set to the implied in BGG, \( \lambda_k \), is estimated, and the indexation to the marginal utility of wealth parameter, λ, is set to 0.

In the Indexation to Rk model there are financial shocks and the elasticity of risk premium, \( \kappa \), is calibrated to 0.041, while the indexation to the return of capital, \( \gamma \), is set to the implied in BGG, \( \gamma_k = \log(1 + \mu) \), where \( \gamma_k = 0.598 \), and the indexation to the marginal utility of wealth parameter, \( \lambda_k \), is set to 0.

In the estimation of the BGG model there are financial shocks and the elasticity of risk premium, \( \kappa \), is calibrated to 0.041, while the indexation to the return of capital, \( \gamma_k \), is estimated, and the indexation to the marginal utility of wealth parameter, \( \lambda_k \), is set to 0.

In the estimation of the BGG and Indexation model there are financial shocks and the elasticity of risk premium, \( \kappa \), is calibrated to 0.041, while the indexation to the return of capital, \( \gamma_k \), and the indexation to the marginal utility of wealth parameter, \( \lambda_k \), are both estimated.

N stands for Norman, B-Beta, G-Gamma, U-Uniform, I-Inv-Gamma- distribution.

Posterior pvalues are from 2 chains of 500,000 draws using a random walk Metropolis algorithm. We discard the initial 250,000 and retain every 5 subsequent draws.
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<th>Output</th>
<th>Monetary policy</th>
<th>Neutral technology</th>
<th>Government</th>
<th>Investment specific technology</th>
<th>Price mark-up</th>
<th>Wage mark-up</th>
<th>Intertemporal preference</th>
<th>Marginal efficiency of investment</th>
<th>Net Worth</th>
<th>Idiosyncratic variance</th>
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<td>2.0</td>
<td>35.9</td>
<td>39.9</td>
<td>7.2</td>
<td>0.0</td>
</tr>
<tr>
<td></td>
<td>R° indexation</td>
<td>0.0</td>
<td>0.0</td>
<td>0.0</td>
<td>0.0</td>
<td>0.0</td>
<td>1.0</td>
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<td>35.9</td>
<td>39.9</td>
<td>7.2</td>
<td>0.0</td>
</tr>
<tr>
<td></td>
<td>R° &amp; λ indexation</td>
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<td>0.0</td>
<td>0.0</td>
<td>0.0</td>
<td>0.0</td>
<td>1.0</td>
<td>2.0</td>
<td>35.9</td>
<td>39.9</td>
<td>7.2</td>
<td>0.0</td>
</tr>
</tbody>
</table>

Table 2: Variance Decomposition at Different Horizons in the JPT, BGG, Indexation to R° and Indexation to R° & λ Models
For the agency cost models (BGG and Indexation) the following parameters are also calibrated: entrepreneurial survival rate, standard deviation of shocks are calibrated to 0.041, while the elasticity of risk premium, $\gamma_k$, is set to 0.5. The indexation to the marginal utility of wealth parameter, $\chi_k$, is set to 0.5.

The indexation to the marginal utility of wealth parameter, $\chi_k$, is estimated, and the indexation to the marginal utility of wealth parameter, $\chi_k$, is set to 0.5.

The indexation to the return of capital parameter, $\nu_k$, is set to 0.041, while the elasticity of risk premium, $\gamma_k$, is calibrated to 0.041.

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REFERENCES


