

Supply Chain Disruption and Reorganization: Theory and Evidence from Ukraine's War

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 - Transmission of negative cost and demand shocks throughout the economy
 - Firms may reorganize production structure and supply chains (e.g., substitution, scaling down)

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 - Transmission of negative cost and demand shocks throughout the economy
 - Firms may reorganize production structure and supply chains (e.g., substitution, scaling down)
- Limited work due to a lack of detailed firm / production network data during wartime & exogenous variation of conflicts

This Paper: 2014 Russia-Ukraine Conflict

- Sudden, intense, but localized conflict in Donbas and annexation of Crimea
- Data: firm-to-firm railroad shipments within Ukraine, 2012–2016
- Reduced-form Evidence:
 - Impacts of supplier & buyer exposure on firms *in nonconflict areas*
 - Outcomes: Firms' output, supplier & buyer links *in nonconflict areas*
- Quantify aggregate effects using a GE model with endogeneous production networks

Preview of Results

- Reduced-form Evidence:
 - \downarrow Relative firm output ($\approx 16\%$), from both supplier and buyer exposure to conflict areas
 - Reorganization of production links away from directly and indirectly exposed firms
 - Supplier exposure $\Rightarrow \uparrow$ number of suppliers & \downarrow buyers *in nonconflict areas*
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- Quantitative GE Model
 - Model sufficient statistics accurately explain observed **firm-level** output loss, with **amplification** from endogenous networks
 - 5.6 % **aggregate** output loss strictly outside conflict areas, with **mitigation** from endogenous networks

Outline

Background and Data

Reduced-Form Evidence

Model

Quantitative Analysis

Conclusion

Background and Data

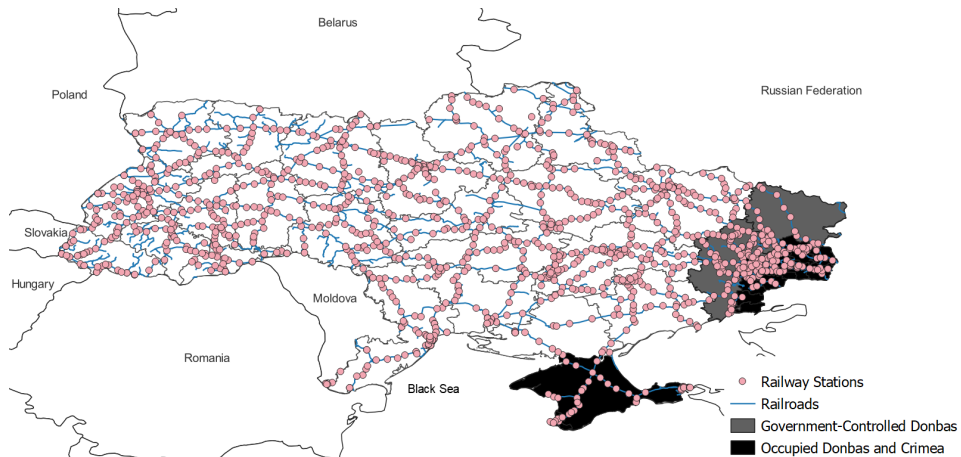
Background: 2014 Russia-Ukraine War

- In February 2014, right after Ukrainian revolution, Russia annexed Crimea and started supporting Donbas separatists
- Sudden, intense, and localized conflict in Donbas regions (until February 2022)
- Donbas (and Crimea) were economic centers of Ukraine before the war
 - Donbas: extractive industry (coal), metallurgy, manufacturing
 - Crimea: agriculture, tourism, some industry
 - Jointly covered 18% of Ukraine's 2013 GDP

Data

- Firm-to-firm railroad shipments within Ukraine, 2012–2016
 - Transactions between ~ 8.5 k firms
 - Sender and receiver firm IDs, dates, weights (kg), freight charges, product codes, origin & destination station codes
 - Focus on inter-firm trade ($\sim 94\%$ of transactions)
 - Impute transaction value using product code (using separate customs data)
- Focusing on railway shipment (vs other shipment modes) unlikely to bias results
 - Railways penetrate all regions in Ukraine, covering 80% of freight in ton-km (Ukr Stat '18)
 - No systematic disruption in railways/roads *outside conflict areas*
 - \Rightarrow *Changes in mode choice outside conflict areas* are likely orthogonal to conflict exposure
- Accounting data for Ukrainian firms, 2010–2018
 - Sources: Spark-Interfax, ORBIS/AMADEUS

Ukrainian Railroads with Stations

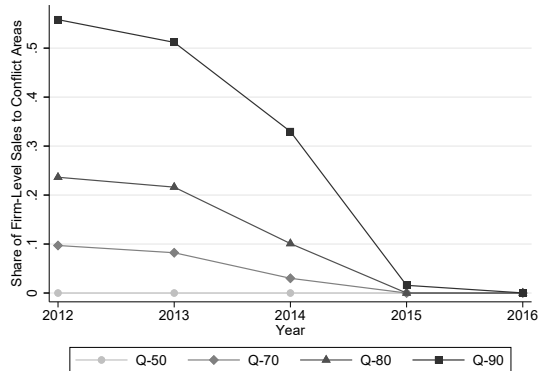
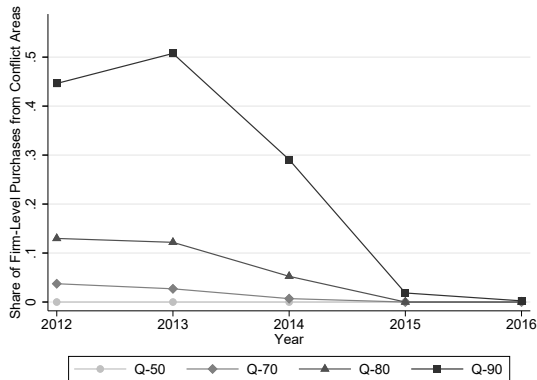


Define “conflict areas” as Crimea and DPR/LPR in Donbas Region hereinafter

Reduced-Form Evidence

Sudden and Large Drop of Trade from & to Conflict Areas

- Weighted fraction of suppliers (left) and buyers (right) from/to conflict areas by firms outside direct conflict areas



event-study

output loss in conflict areas

trade exposure to Russia

Firm-Level Impacts of Conflict Exposure

Difference-in-differences specification:

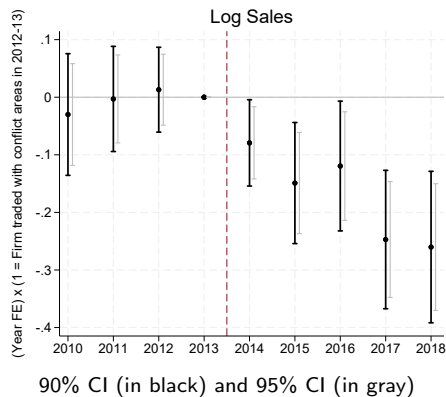
$$Y_{ft} = \gamma \times Post_t \times SupplierExposure_f + \beta \times Post_t \times BuyerExposure_f + \alpha_f + \delta_t + \varepsilon_{ft}$$

- f : firms outside conflict areas
- Y_{ft} : sales, linkages outside conflict areas
- $SupplierExposure_f$: Value share of shipment *from* conflict areas in 2012-13
- $BuyerExposure_f$: Value share of shipment *to* conflict areas in 2012-13

No pretrends, robust to region-time FE, industry-time FE, and trade with Russia controls

Large Negative Impacts of Conflict Exposure on Sales

$$\log Sales_{ft} = \gamma_t \times \mathbb{1}[\text{TradeConflictExposure}_f > 0] + \alpha_f + \delta_t + \varepsilon_{ft}$$



Impacts of Supplier and Buyer Conflict Exposure on Sales

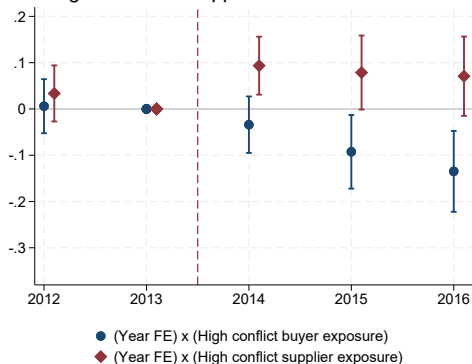
| | (1) Log Sales | (2) No Sales Reported | (3) Log Sales | (4) No Sales Reported | (5) Log Sales | (6) No Sales Reported |
|--|----------------------|-----------------------------|----------------------|-----------------------------|----------------------|-----------------------------|
| Post-2014 × 1[Firm traded with conflict areas, 2012–13] | -0.162*** (0.046) | 0.070*** (0.010) | | | | |
| Post-2014 × Firm's buyer conflict exposure, 2012–13 | | | -0.215** (0.100) | 0.060*** (0.023) | | |
| Post-2014 × Firm's seller conflict exposure, 2012–13 | | | -0.280*** (0.100) | 0.066*** (0.022) | | |
| Post-2014 × 1[High firm's buyer conflict exposure, 2012–13] | | | | | -0.190*** (0.058) | 0.058*** (0.012) |
| Post-2014 × 1[High firm's seller conflict exposure, 2012–13] | | | | | -0.139** (0.054) | 0.043*** (0.012) |
| Firm FE | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Year FE | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Mean | 16.899 | 0.291 | 16.899 | 0.291 | 16.899 | 0.291 |
| SD | 2.482 | 0.454 | 2.482 | 0.454 | 2.482 | 0.454 |
| Observations | 35,439 | 50,202 | 35,439 | 50,202 | 35,439 | 50,202 |
| Number of Firms | 4,775 | 5,578 | 4,775 | 5,578 | 4,775 | 5,578 |

robustness

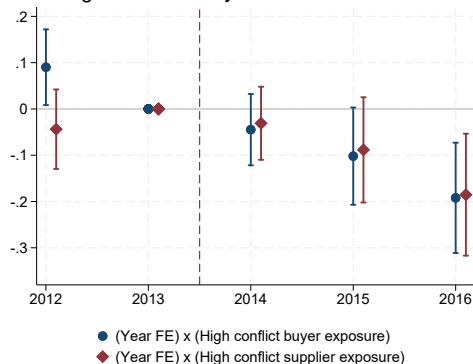
Reorganization of Supplier and Buyer Linkages Outside Conflict Areas table

$$Y_{ft} = \gamma_t \times \mathbb{1}[\text{HighSupplierExposure}_f] + \beta_t \times \mathbb{1}[\text{HighBuyerExposure}_f] + \alpha_f + \delta_t + \varepsilon_{ft}$$

Log Number of *Suppliers* in Nonconflict Areas



Log Number of *Buyers* in Nonconflict Areas



- **Supplier exposure:** substitute suppliers toward nonconflict areas, but lose their buyers
- **Buyer exposure:** reduce input demand, leading to losing buyers even in nonconflict areas

Model

- Regions: $i \in \mathcal{L}$
- Measure L_i of HHs in region i , supply one unit of labor at competitive wage w_i
- Heterogeneous firm types in region i : $\omega \in \Omega_i$, measure $N_i(\omega)$
 - e.g., heterogeneity in prior connection to conflict areas
- Transactions can occur as long as they are connected by (endogeneous) networks
 - Firms are identical within types \Rightarrow measure of supplier linkages across types summarize the network architecture
 - Iceberg costs across locations, sectors, and types
- Single sector for presentation

Technology and Trade Flows

- Firm type $\omega \in \Omega_i$'s production technology

$$Y_i(\omega) = Z_i(\omega) \left(\frac{L_i(\omega)}{\beta} \right)^\beta \left(\frac{Q_i(\omega)}{1-\beta} \right)^{1-\beta}, \quad Q_i(\omega) = \left(\sum_{u \in \mathcal{L}} \sum_{v \in \Omega_u} M_{ui}(v, \omega) q_{ui}(v, \omega)^{\frac{\sigma-1}{\sigma}} \right)^{\frac{\sigma}{\sigma-1}}$$

- $M_{ui}(v, \omega)$: measure of supplier linkages for firm $\omega \in \Omega_u$ with suppliers $v \in \Omega_i$ (endogeneous)
- Supplier linkages benefit production through love-of-variety in CES
- Continuum of connections \Rightarrow constant markup $1/\sigma$
- Nominal trade flows:

$$X_{ui}(v, \omega) = M_{ui}(v, \omega) \tau_{ui}(v, \omega)^{1-\sigma} C_u(v)^{1-\sigma} D_i(\omega)$$

- Intermediate goods sales by firm type ω :

$$R_i(\omega) = Z_i(\omega)^{\sigma-1} w_i(\omega)^{\beta_L(1-\sigma)} \underbrace{\mathcal{A}_i^S(\omega)}_{\text{supplier access}} \underbrace{\mathcal{A}_i^B(\omega)}_{\text{buyer access}}$$

$$\mathcal{A}_i^S(\omega) \equiv \left(\sum_{u \in \mathcal{L}} \sum_{v \in \Omega_u} M_{ui}(v, \omega) \tau_{ui}(v, \omega)^{1-\sigma} C_u(v)^{1-\sigma} \right)^{\beta}$$

$$\mathcal{A}_i^B(\omega) \equiv \sum_{d \in \mathcal{L}} \sum_{\psi \in \Omega_d} M_{id}(\omega, \psi) \tau_{id}(\omega, \psi)^{1-\sigma} D_d^*(\psi)$$

- Summarize four variables that shape firm-level output
- Use this expression to assess what drives firm-level output decline empirically

- Equilibrium measure of supplier connections are given by:

$$M_{ui}(v, \omega) = \underbrace{K_{ui}(v, \omega)}_{\text{exog factor}} \frac{X_{ui}(v, \omega)^{\lambda^S + \lambda^B}}{e_u(v)^{\lambda^S} e_i(\omega)^{\lambda^B}}, \quad \underbrace{e_i(\omega)}_{\text{link formation cost}} = w_i(\omega)^\mu C_i(\omega)^{1-\mu}$$

- Can be microfounded through search & matching (Boehm & Oberfield '23; Demir et al 24; Arkolakis et al '24) or entry (Melitz & Redding '14)
- Households with CD-CES preferences purchase final goods from local firms
- Labor, intermediate goods, final goods markets clear

Quantitative Analysis

- 25 oblasts (provinces) + 3 “conflict area” (DPR, LPR, Crimea)
- Three sectors: mining, manufacturing, other
- 4 firm types within region-sector based on high/low supplier and buyer exposure (80th percentiles) prior to the conflict
- Trade flows and production linkages: from railway shipment data
- Parameters: detail
 - $\{\beta_{L,m}, \beta_{km}, \alpha_k\}$: from IO table
 - $\{\sigma_k\}$: from profit to revenue ratio
 - $\{\lambda^S, \lambda^B, \mu\}$: target network reorganization in response to conflict exposure ($\lambda^S = \lambda^B = 0.15, \mu = 1$)

Assessing the Mechanism Behind Firm-level Output Reduction

- Model implies

$$\log R_{i,m,t}(\omega) = \log \left[w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \mathcal{A}_{i,m,t}^S(\omega) \mathcal{A}_{i,m,t}^B(\omega) \right] + \log Z_{i,m,t}(\omega)^{\sigma_m-1}$$

- We estimate:

$$\log R_{i,m,t}(\omega) = \gamma \log \left[w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega) \right] + \eta_{i,m}(\omega) + \nu_{i,t} + \delta_{m,t} + \epsilon_{i,m,t}(\omega)$$

- $\tilde{\mathcal{A}}_{i,t}^S(\omega)$, $\tilde{\mathcal{A}}_{i,t}^B(\omega)$: estimate from panel gravity equations using railway data detail
- IV: high conflict supplier and buyer exposure \times post
- Test $\gamma = 1$: conflict exposure affects $R_{i,t}(\omega)$ through wages & access, not through unobserved TFP changes

Model Sufficient Statistics Accurately Explain Firm-Level Output Changes

| | $\log R_{i,m,t}(\omega)$ | | |
|--|--------------------------|----------------|----------------|
| | (1) | (2) | (3) |
| Panel A: With Link Adjustment | | | |
| $\log w_{i,t}^{\beta_{m,t}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 0.85 (0.12) | 0.88 (0.13) | 0.83 (0.11) |
| p-value (coefficient = 1) | 0.23 | 0.35 | 0.13 |
| Effective First-Stage F-Statistics | 45.7 | 43.1 | 49 |
| Panel B: No Link Adjustment | | | |
| $\log w_{i,t}^{\beta_{m,t}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 1.61 (0.36) | 1.72 (0.41) | 1.71 (0.37) |
| p-value (coefficient = 1) | 0.09 | 0.08 | 0.06 |
| Effective First-Stage F-Statistics | 16.3 | 14.7 | 16.3 |
| Firm-Type-Region-Sector Fixed Effects | X | X | X |
| Year Fixed Effects | X | X | X |
| Sector \times Year Fixed Effects | | X | X |
| Region \times Year Fixed Effects | | | X |
| Observations | 434 | 434 | 434 |

- Cost & demand effects, not TFP changes, explain firm-level output decline
- $\gamma > 1$ in Panel B \Rightarrow model with no link adjustment underpredicts sales reduction

shut down only supplier links

only buyer

use all years

gravity with agg. flows

Quantify Aggregate Effects Outside Conflict Areas

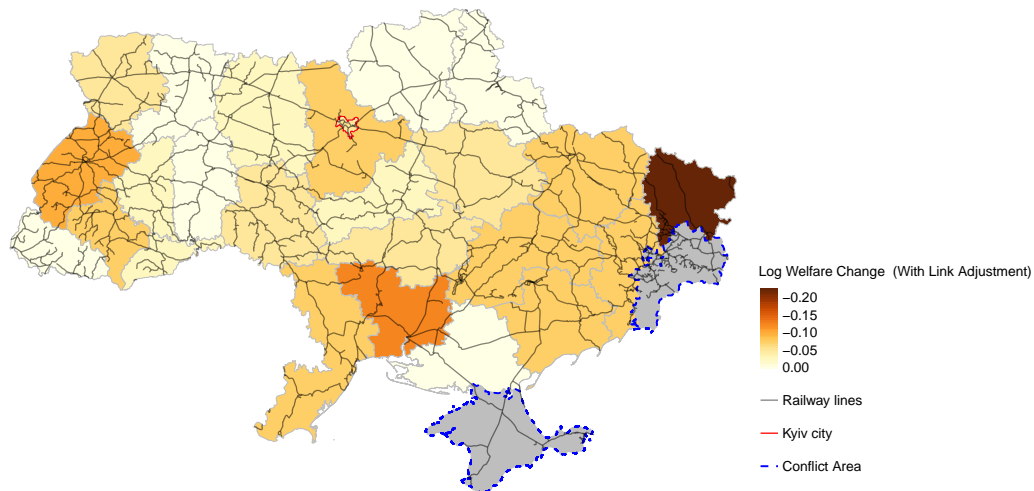
- Cost/demand propagation accurately summarizes (relative) firm-level output decline, network reorganization *amplifies* this effect
- What about aggregate effects?
- Calibrate model with 2013 trade and production linkage patterns, simulate $\tau_{ui,km}(v, \omega) \rightarrow \infty$ if u or i is in conflict areas

Large Aggregate Output Loss in Nonconflict Areas, **Mitigated** by Reorganization

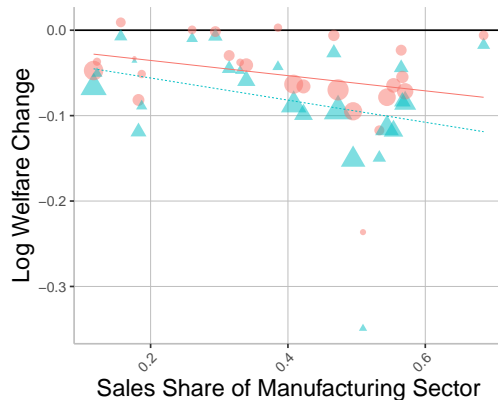
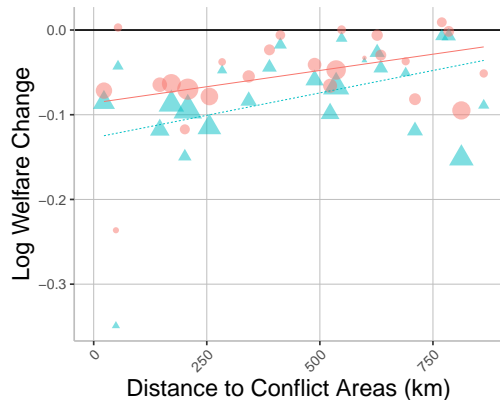
| Real GRP Changes (%) | Mean | 25%-ile | 50%-ile | 75%-ile |
|--|------|---------|---------|---------|
| (1) With Link Adjustment | -5.6 | -7.2 | -6.3 | -3.3 |
| (2) No Link Adjustment | -8.4 | -11.5 | -8.6 | -4.5 |
| (3) With Link Adjustment (Shock to DPR) | -1.8 | -2.2 | -1.3 | -0.4 |
| (4) With Link Adjustment (Shock to LPR) | -2.6 | -4.1 | -2.4 | -1.6 |
| (5) With Link Adjustment (Shock to Crimea) | -0.9 | -1 | -0.3 | 0.1 |

- Large aggregate welfare loss, *mitigated* by reorganization
- Coordinated shocks to DPR, LPR, Crimea have slight additional cost than cumulative effects from independent shocks (5.3% vs 5.6%) robustness

Negative Welfare Effects Even for Distant Region from Conflict Areas



Welfare Effects By Distance to Conflict Areas and Manufacturing Share



Conclusion

Conclusion

- Provide reduced-form evidence of significant supply chain disruption and reorganization during 2014 Ukraine War, beyond Donbas and Crimea
- Supply chain reorganization **amplifies** firm-level output loss but **mitigate** aggregate output loss
- Highlights a key mechanism in which localized conflict often have far-reaching detrimental consequences for the broader economy (Rohner & Thoenig '21)



Appendix



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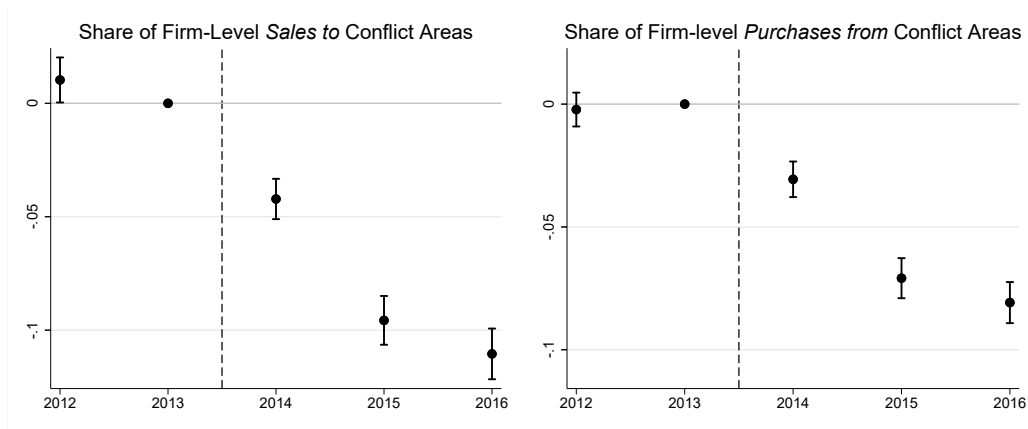
Tesla to Halt Production in Germany as Red Sea Conflict Hits Supply Chains

Disruption related to attacks on ships by Houthi rebels raise risk of supply-chain crisis in Europe

By [William Boston](#) [Follow](#), [Costas Paris](#) [Follow](#) and [Benoit Faucon](#) [Follow](#)

Updated Jan. 12, 2024 at 1:45 pm ET

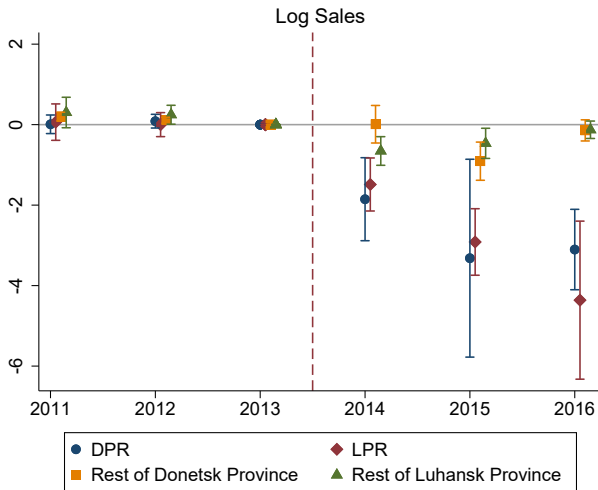
Sudden and Large Drop of Trade from & to Conflict Areas [go back](#)



Sudden and Large Drop of Aggregate Firm Sales in Conflict Areas [go back](#)

$$\begin{aligned} Y_{rt} = & \beta_t^{LPR} \times LPR_r \times Post_t \\ & + \beta_t^{DPR} \times DPR_r \times Post_t \\ & + \beta_t^{DON} \times Donetsk_r \times Post_t \\ & + \beta_t^{LUH} \times Luhansk_r \times Post_t \\ & + \alpha_r + \kappa_t + \varepsilon_{rt} \end{aligned}$$

- r : rayon (district)
- Exclude Crimea due to data quality after the annexation
- Consistent with decline in nighttime light (Kochnev '19)



Summary Statistics of exposure with Conflict Areas and with Russia [go back](#)

| | Observations | Mean | SD | Min | Max |
|--|--------------|------|------|-----|-----|
| 1[Firm traded with conflict areas, 2012–13] | 50,202 | 0.55 | 0.50 | 0 | 1 |
| Firm's buyer conflict exposure, 2012–2013 | 50,202 | 0.09 | 0.22 | 0 | 1 |
| Firm's supplier conflict exposure, 2012–2013 | 50,202 | 0.10 | 0.23 | 0 | 1 |
| 1[High firm's buyer conflict exposure, 2012–13] | 50,202 | 0.19 | 0.39 | 0 | 1 |
| 1[High firm's supplier conflict exposure, 2012–13] | 50,202 | 0.19 | 0.39 | 0 | 1 |
| 1[Firm traded with Russia in 2012–2013] | 50,202 | 0.24 | 0.43 | 0 | 1 |

Impacts of Supplier and Buyer Conflict exposure on Sales: Robustness [go back](#)

| | (1) Baseline | (2) Strictly balanced panel | (3) Latitude & longitude | (4) Latitude & longitude | (5) Distance to conflict areas | (6) Distance to conflict areas | (7) 2-digit industry × post | (8) Region FE × post | (9) Pre-conflict trade with Russia | (10) Pre-conflict trade partners | (11) Omitting Donetsk oblast | (12) Omitting Luhansk oblast | (13) Omitting Kyiv |
|--|----------------------|--------------------------------------|--------------------------------|--------------------------------|--------------------------------------|--------------------------------------|--------------------------------------|----------------------------|---|---|---------------------------------------|---------------------------------------|--------------------------|
| Post-2014 × 1[Firm traded with conflict areas, 2012–13] | -0.162*** (0.046) | -0.100** (0.045) | -0.139*** (0.046) | -0.130*** (0.046) | -0.141*** (0.046) | -0.146*** (0.046) | -0.110** (0.047) | -0.125*** (0.046) | -0.149*** (0.046) | -0.133*** (0.046) | -0.159*** (0.046) | -0.126*** (0.047) | |
| Post-2014 × Latitude | | | 0.061*** (0.016) | -1.251 (0.923) | | | | | | | | | |
| Post-2014 × Longitude | | | -0.020*** (0.005) | -1.055*** (0.290) | | | | | | | | | |
| Post-2014 × Latitude ² | | | | 0.006 (0.009) | | | | | | | | | |
| Post-2014 × Longitude ² | | | | -0.002 (0.001) | | | | | | | | | |
| Post-2014 × Latitude × longitude | | | | 0.023*** (0.006) | | | | | | | | | |
| Post-2014 × Distance to conflict area | | | | | 0.505*** (0.098) | | | | | | | | |
| Post-2014 × Distance to LPR or DPR | | | | | | 0.388*** (0.079) | | | | | | | |
| Post-2014 × 1[Firm imported from Russia, 2012–13] | | | | | | | | -0.218*** (0.060) | | | | | |
| Post-2014 × 1[Firm exported to Russia, 2012–13] | | | | | | | | -0.224*** (0.061) | | | | | |
| Post-2014 × # of preconflict trade partners | | | | | | | | | -0.000* (0.000) | | | | |
| Firm FE | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | |
| Year FE | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | |
| Mean | 16.899 | 17.237 | 16.900 | 16.900 | 16.900 | 16.900 | 16.934 | 16.899 | 16.899 | 16.857 | 16.901 | 16.847 | |
| SD | 2.482 | 2.291 | 2.481 | 2.481 | 2.481 | 2.481 | 2.473 | 2.482 | 2.482 | 2.455 | 2.479 | 2.435 | |
| Observations | 35,439 | 24,273 | 35,334 | 35,334 | 35,334 | 35,334 | 33,812 | 35,439 | 35,439 | 33,640 | 34,888 | 30,383 | |
| Number of Firms | 4,775 | 2,697 | 4,753 | 4,753 | 4,753 | 4,753 | 4,558 | 4,775 | 4,775 | 4,530 | 4,700 | 4,007 | |

Impacts of Supplier and Buyer Conflict Exposure on Linkages [go back](#)

| | (1) Log # of Suppliers in Nonconflict Areas | (2) Log # of Buyers in Nonconflict Areas | (3) Log # of Suppliers in Nonconflict Areas | (4) Log # of Buyers in Nonconflict Areas |
|--|---|--|---|--|
| Post-2014 × Firm's buyer conflict exposure, 2012–13 | -0.071 (0.061) | -0.156 (0.100) | | |
| Post-2014 × Firm's seller conflict exposure, 2012–13 | 0.263*** (0.068) | -0.203** (0.100) | | |
| Post-2014 × 1[High firm's buyer conflict exposure, 2012–13] | | | -0.089*** (0.033) | -0.156*** (0.043) |
| Post-2014 × 1[High firm's seller conflict exposure, 2012–13] | | | 0.064** (0.032) | -0.077* (0.046) |
| Firm FE | ✓ | ✓ | ✓ | ✓ |
| Year FE | ✓ | ✓ | ✓ | ✓ |
| Mean | 1.790 | 1.945 | 1.790 | 1.945 |
| SD | 1.243 | 1.495 | 1.243 | 1.495 |
| Observations | 18,390 | 11,881 | 18,390 | 11,881 |
| Number of Firms | 4,281 | 3,031 | 4,281 | 3,031 |

Multi-Sector Model

- Firms belong to a sector $k \in K$
- Cobb-Douglas production with input share β_{km} with sector-specific elasticity of substitution σ_k

$$Y_{i,m}(\omega) = Z_{i,m}(\omega) \left(\frac{L_{i,m}(\omega)}{\beta_{m,L}} \right)^{\beta_{m,L}} \prod_{k \in K} \left(\frac{Q_{i,km}(\omega)}{\beta_{km}} \right)^{\beta_{km}}$$

$$Q_{i,km}(\omega) = \left(\sum_{u \in \mathcal{L}} \sum_{v \in \Omega_{u,k}} M_{ui,km}(v, \omega) q_{ui,km}(v, \omega)^{\frac{\sigma_k - 1}{\sigma_k}} \right)^{\frac{\sigma_k}{\sigma_k - 1}}$$

- Final consumption share α_k
- Measure of linkages: $M_{ui,km}(v, \omega)$

Calibrate Structural Parameters from Ukraine's Pre-War IO Table [go back](#)

- $\{\beta_{L,m}, \beta_{km}, \alpha_k\}$: Input and final expenditure shares
- $\{\sigma_k\}$: Pre-tax profit to revenue ratio

| | Sectors (m) | | |
|----------------------------|-----------------|---------------|-------|
| | Mining | Manufacturing | Other |
| (a) β_{km} | | | |
| $k = \text{Mining}$ | 0.11 | 0.12 | 0.06 |
| $k = \text{Manufacturing}$ | 0.18 | 0.33 | 0.18 |
| $k = \text{Other}$ | 0.36 | 0.45 | 0.40 |
| (b) $\beta_{m,L}$ | 0.35 | 0.10 | 0.36 |
| (c) α_m | 0.01 | 0.60 | 0.39 |
| (d) σ_m | 4.8 | 8.1 | 5.0 |

Estimating Supplier and Buyer Accesses [go back](#)

- Model-predicted trade flows (with time subscript t):

$$\frac{X_{ui,t}(v, \omega)}{M_{ui,t}(v, \omega)} = C_{u,t}(v)^{1-\sigma} D_{i,t}(\omega) \tau_{ui,t}(v, \omega)^{1-\sigma}$$

- We estimate a three-way fixed-effect model by PPML:

$$\frac{X_{ui,t}(v, \omega)}{M_{ui,t}(v, \omega)} = \xi_{u,t}(v) \zeta_{i,t}(\omega) \eta_{ui}(v, \omega) \epsilon_{ui,t}(v, \omega)$$

- Using these estimates,

$$\tilde{\mathcal{A}}_{i,t}^S(\omega) = \left(\sum_{u \in \mathcal{L}} \sum_{v \in \Omega_u} M_{ui,t}(v, \omega) \tilde{\eta}_{ui}(v, \omega) \tilde{\xi}_{u,t}(v) \right)^{1-\beta}$$

$$\tilde{\mathcal{A}}_{i,t}^B(\omega) = \sum_{d \in \mathcal{L}} \sum_{\psi \in \Omega_d} M_{id,t}(\omega, \psi) \tilde{\eta}_{ui}(\omega, \psi) \tilde{\zeta}_{i,t}(\psi)$$

Market clearing (multiple sector)

- Final goods sales

$$R_{i,m}^F(\omega) = \frac{\varsigma_m N_{i,m}(\omega) C_{i,m}(\omega)^{1-\sigma_k}}{\left(P_{i,m}^F\right)^{1-\sigma_m}} \alpha_m E_i L_i$$

- Intermediate goods sales

$$R_{i,m}(\omega) = \tilde{\varsigma}_m Z_{i,m}(\omega)^{\sigma_m-1} w_i^{\beta_{m,L}(1-\sigma_m)} \mathcal{A}_{i,m}^S(\omega) \mathcal{A}_{i,m}^B(\omega),$$

- Labor market clearing

$$w_i L_i = \sum_{m \in K} \beta_{L,m} \frac{\sigma_m - 1}{\sigma_m} \left(R_{i,m}(\omega) + R_{i,m}^F(\omega) \right),$$

- Firm profit

$$\pi_{i,m}(\omega) = \sum_{m \in K} \frac{1}{\sigma_m} \left(R_{i,m}(\omega) + R_{i,m}^F(\omega) \right).$$

Model Validation: Shut Down Only Buyer Linkage Changes [go back](#)

| | $\log R_{i,m,t}(\omega)$ | | |
|--|--------------------------|----------------|----------------|
| | (1) | (2) | (3) |
| $\log w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 3.49 (1.67) | 4.44 (2.60) | 4.04 (1.74) |
| p -value (coefficient = 1) | 0.13 | 0.19 | 0.08 |
| Effective First-Stage F-Statistics | 5 | 3.2 | 5.7 |
| Firm-Type-Region-Sector Fixed Effects | X | X | X |
| Year Fixed Effects | X | X | X |
| Sector \times Year Fixed Effects | | X | X |
| Region \times Year Fixed Effects | | | X |
| Observations | 433 | 433 | 433 |
| Adjusted R^2 | -0.29 | -0.81 | -0.19 |

Model Validation: Shut Down Only Supplier Linkage Changes [go back](#)

| | $\log R_{i,m,t}(\omega)$ | | |
|--|--------------------------|----------------|----------------|
| | (1) | (2) | (3) |
| $\log w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 1.19 (0.17) | 1.20 (0.17) | 1.10 (0.13) |
| p-value (coefficient = 1) | 0.26 | 0.24 | 0.43 |
| Effective First-Stage F-Statistics | 37.2 | 38.2 | 56.9 |
| Firm-Type-Region-Sector Fixed Effects | X | X | X |
| Year Fixed Effects | X | X | X |
| Sector \times Year Fixed Effects | | X | X |
| Region \times Year Fixed Effects | | | X |
| Observations | 438 | 438 | 438 |
| Adjusted R ² | 0.86 | 0.88 | 0.92 |

Model Validation: Use All Years [go back](#)

| | $\log R_{i,m,t}(\omega)$ | | |
|--|--------------------------|----------------|----------------|
| | (1) | (2) | (3) |
| $\log w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 0.77 (0.12) | 0.78 (0.12) | 0.71 (0.10) |
| p -value (coefficient = 1) | 0.05 | 0.08 | 0.00 |
| Effective First-Stage F-Statistics | 43.4 | 42.6 | 55.8 |
| Firm-Type-Region-Sector Fixed Effects | X | X | X |
| Year Fixed Effects | X | X | X |
| Sector \times Year Fixed Effects | | X | X |
| Region \times Year Fixed Effects | | | X |
| Observations | 1,085 | 1,085 | 1,085 |
| Adjusted R^2 | 0.90 | 0.90 | 0.92 |

Model Validation: Estimate Gravity using Aggregate Flows [go back](#)

| | $\log R_{i,m,t}(\omega)$ | | |
|--|--------------------------|----------------|----------------|
| | (1) | (2) | (3) |
| $\log w_{i,t}^{\beta_{m,L}(1-\sigma_m)} \tilde{\mathcal{A}}_{i,m,t}^S(\omega) \tilde{\mathcal{A}}_{i,m,t}^B(\omega)$ | 1.61 (0.36) | 1.72 (0.41) | 1.71 (0.37) |
| p-value (coefficient = 1) | 0.09 | 0.08 | 0.06 |
| Effective First-Stage F-Statistics | 16.3 | 14.7 | 16.3 |
| Firm-Type-Region-Sector Fixed Effects | X | X | X |
| Year Fixed Effects | X | X | X |
| Sector \times Year Fixed Effects | | X | X |
| Region \times Year Fixed Effects | | | X |
| Observations | 434 | 434 | 434 |
| Adjusted R ² | 0.69 | 0.65 | 0.69 |

Counterfactual Simulation: Robustness [go back](#)

| Alternative Specifications | λ^S | λ^B | μ | Average Welfare Change (Baseline) | Average Welfare Change (No Link Adjustment) |
|--|-------------|-------------|-------|--------------------------------------|--|
| (a) Baseline | 0.15 | 0.15 | 1.00 | -5.6 | -8.4 |
| (b) Set $\lambda^B = 0$ | 0.30 | 0.00 | 1.00 | -5.5 | -8.4 |
| (c) Set $\lambda^S = 0$ | 0.00 | 0.30 | 1.00 | -5.6 | -8.4 |
| (d) Set $\mu = 0$ | 0.15 | 0.15 | 0.00 | -6.6 | -8.5 |
| (e) Set $\delta_m = 0.5$ | 0.15 | 0.15 | 1.00 | -5.6 | -8.4 |
| (f) Define Types by Link Exposures | 0.15 | 0.15 | 1.00 | -5.9 | -9.1 |
| (g) Define Types by Weight Exposures | 0.15 | 0.15 | 1.00 | -5.6 | -8.2 |
| (h) Define Types by Exposure and Firm Size | 0.15 | 0.15 | 1.00 | -6.6 | -9.9 |